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Abstract	Identifying gene–gene and gene–environment interactions may help us to better describe the genetic architecture for complex traits. While advances have been made in identifying genetic variants associated with complex traits through more dense panels of genetic variants and larger sample sizes, genome-wide interaction analyses are still limited in power to detect interactions with small effect sizes, rare frequencies, and higher order interactions. This chapter outlines methods for detecting both gene-gene and gene-environment interactions both through explicit tests for interactions (i.e., ones in which the interaction is tested directly) and non-explicit tests (i.e., ones in which an interaction is allowed for in the test, but does not test for the interaction directly) as well as approaches for increasing power by reducing the search space. Issues relating to multiple test correction, replication, and the reporting of interaction result in publications.
Keywords (separated by ‘-’)	Interaction - Epistasis - Environment - GWAS - Power - Replication

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Abstract 4

Identifying gene–gene and gene–environment interactions may help us to better describe the genetic architecture for complex traits. While advances have been made in identifying genetic variants associated with complex traits through more dense panels of genetic variants and larger sample sizes, genome-wide interaction analyses are still limited in power to detect interactions with small effect sizes, rare frequencies, and higher order interactions. This chapter outlines methods for detecting both gene-gene and gene-environment interactions both through explicit tests for interactions (i.e., ones in which the interaction is tested directly) and non-explicit tests (i.e., ones in which an interaction is allowed for in the test, but does not test for the interaction directly) as well as approaches for increasing power by reducing the search space. Issues relating to multiple test correction, replication, and the reporting of interaction result in publications.

Key words Interaction, Epistasis, Environment, GWAS, Power, Replication 15

1 Introduction 16

An interaction results when the effect of one factor is only evident in the presence of another. These factors could be genetic markers and/or environmental exposures. Much has been written on the topic of gene-gene (also known as epistasis) and gene-environment interactions with several comprehensive reviews of study designs and methods for analyzing both gene-gene [1–3] and gene-environment interactions [4, 5]. But why are we interested in studying interactions? Both gene-gene and gene-environment and potentially gene-gene-environment interactions allow us to better describe the underlying genetic architecture of a particular trait and as such we can begin to fill in the missing heritability [6] for a particular phenotype.

Biological interactions were originally defined as the situation when the phenotypic effect of one gene was only evident in the presence of a second gene [7]. In contrast, a statistical interaction is defined as a departure from a linear model combining two or more genetic factors (or a genetic factor and environmental factor) [8].

Another way to think about this is that biological interactions are 34
observed at the individual level and statistical interactions at the 35
population level, but this does not imply that observing evidence of 36
one will lead to observing evidence of another [9]. The traditional 37
method to test for statistical interactions is to use a regression-based 38
model that includes main effect variables representing each genetic 39
factor and an interaction term (the product of the main effect 40
variables), and then testing for the significance of the interaction 41
term after adjusting for the main effects. However, several other 42
methods exist to test for statistical interactions and these will be 43
discussed as well. 44

We now have the ability to conduct very dense genome-wide 45
association studies with up to five million genetic markers geno- 46
typed at one time and millions more imputed using large reference 47
panels from population-based sequencing projects. The combina- 48
torial problem can be immense if you attempt to look at all possible 49
marker combinations with little power to detect significant inter- 50
actions after accounting for all interactions tested. Given the expo- 51
nential improvements in computational power and the relative ease 52
of parallel computing, the computational hurdles of examining all 53
pairwise gene-gene interactions are not insurmountable, and 54
exhaustive searches of higher order interactions will follow. How- 55
ever, the immense power constraints on currently available sample 56
sizes, while improving with studies routinely examining 10,000 to 57
more than 100,000 subjects, are still underpowered to detect inter- 58
actions with modest effect sizes ($OR < 1.2$) and low frequency 59
variants (minor allele frequency [MAF] < 0.01). Though marker 60
pruning using linkage disequilibrium can reduce this problem 61
somewhat, there are several other data reduction approaches that 62
will be discussed based on prior biological knowledge or statistical 63
evidence for association. These can mitigate the multiple testing 64
burden but depend heavily on the quality of this prior information. 65

Similar to gene-gene interactions, gene-environment interac- 66
tions arise when the effect of a genetic factor on a phenotype is 67
dependent on the presence or absence of an environmental factor. 68
Statistically, this can be tested in a similar fashion as for gene-gene 69
interactions traditionally done using a regression-based model with 70
main effects for the genetic factor and environmental factor and an 71
interaction term and testing for the significance of the interaction 72
term. This environmental factor may be one traditionally thought 73
of as an environmental exposure such as smoking, indoor NO_2 74
levels, or sun exposure, but these could also be other potentially 75
genetically influenced phenotypes such as obesity, blood glucose 76
levels, or birthweight that may be influencing the effect of the 77
genetic factor depending on the value of these secondary 78
phenotypes. 79

As outlined below, direct assessment of interactions (termed as 80
“explicit” test) is often less powerful than joint tests that include 81

either an interaction term or allow for interaction but do not test 82
for the significance of the interaction directly (termed as “non- 83
explicit” test). This classification of tests for interactions will be 84
used throughout this chapter as different approaches and tests are 85
discussed. 86

2 Materials 87

2.1 Data 88

2.1.1 Genotype Data 89

The core of both gene-gene and gene-environment interaction 88
analyses is obtaining high-quality genotype data. One may use 89
data generated specifically for a particular project, but there are 90
also many outstanding datasets available for analysis from a number 91
of online repositories. 92

If samples are being genotyped for the specific project there are 93
many options available. There are whole genome microarray panels 94
of markers based on relatively even coverage across the genome, 95
panels of markers that maximize the coverage for specific race/ 96
ethnic groups, and panels that allow custom markers to be added- 97
on to existing panels to increase coverage of specific genes of 98
interest or previously associated markers, for example. There are 99
also panels of markers that target specific regions of the genome 100
such as the exonic regions, cancer-associated genes, and metabolic 101
genes. The selection of the right microarray will depend on your 102
specific study hypotheses, type of study, and budget. Lower 103
throughput genotyping can be done for single variants to hundreds 104
of markers simultaneously depending again on the hypotheses and 105
goals of the project. 106

Genotyping known panels of markers is not the only choice for 107
interaction studies. One could choose to utilize sequencing-based 108
approaches to genotype unknown variants and/or low frequency 109
and rare variants. Whole genome or whole exome sequencing could 110
be utilized for hypothesis-free analyses or targeted sequencing if 111
specific genes or pathways are hypothesized to be involved in the 112
interaction(s). 113

If secondary data analysis is an option, there are a plethora of 114
datasets with genome-wide data available as well as extensive phe- 115
notype data. Two such repositories are the database of genotype and 116
phenotype (dbGaP) maintained by NCBI primarily of studies con- 117
ducted in the United States (<https://www.ncbi.nlm.nih.gov/gap>) 118
and the European Genome-phenome Archive (EGA) maintained 119
by EBI primarily of studies conducted in Europe ([https://www.ebi.](https://www.ebi.ac.uk/ega/home) 120
[ac.uk/ega/home](https://www.ebi.ac.uk/ega/home)). These databases contain hundreds of datasets 121
accessible through an application to the respective data access 122
committees. 123

2.1.2 Environmental Data 124

The environmental factors that could be considered in a study of 125
gene-environment interactions are extremely broad. These include 126

chemical (e.g., polychlorinated biphenyl (PCB)), physical (e.g., airborne particulate matter), biological (e.g., viral infections), and lifestyle (e.g., physical activity). The measurement of each of these individual environmental factors is going to vary widely depending on the environmental factor of interest. This could range from measure NO₂ levels in the air via chemiluminescence, measuring radiation exposure via a dosimeter, conducting a daily food diary to estimate saturated fat intake, or reviewing charts to collect data on BMI history. The discussion in this chapter will focus on environmental exposures at one time point, but there is some evidence that longitudinal environmental data may increase the power to detect gene-environment interactions for common diseases [10].

2.1.3 Biological/ Functional Data

For analyses that are pursuing a hypothesis-driven approach and/or filtering based on biological information, one may want to utilize prior biological or functional data. There are a variety of databases and programs that can be accessed to provide this type of information. One could simply curate information from publications in the scientific literature through systemic reviews of publications in databases such as PubMed. Data on functionality of variants can be obtained more systematically from databases such as ENCODE (Encyclopedia of DNA Elements) [11] that contains a comprehensive list of functional elements at both the RNA and protein levels and is available for viewing or downloading from the UCSC Genome Browser (www.genome.ucsc.edu/ENCODE). Direct annotation of variants could be conducted using a program such as Annovar [12] to annotate variants as to their respective genes, coding vs. noncoding, and predicted functional consequence. An alternative annotation program more directly related to filtering variants for interaction analyses is Biofilter which allows for the annotation of variants based on previous association studies and biological knowledge, filtering variants based on specific biological hypotheses, and building sets of testable variant interactions based on implication indices compiled from available data [13].

2.1.4 Previous Statistical Data

For analyses filtering variants based on prior statistical knowledge, data from one's own GWAS or single-variant association study could be used, results mined from previous publications or, alternatively, association results obtained from databases such as the GWAS Catalog (<http://www.ebi.ac.uk/gwas/>). While results from previous publications or the GWAS Catalog are a convenient and useful resource, they have the disadvantage of being biased toward reporting only genome-wide significant results and other nominally significant results will likely not be available and should be kept in mind when planning the analysis approach.

2.2 Software

Below are a listing of programs that can be used to conduct gene-gene and/or gene-environment interaction analyses with other

programs mentioned and described throughout the chapter. This is not an exhaustive list and is only an example of programs that are commonly used to assess interactions in genetic epidemiological studies. Additional programs not directly related to the interaction analysis such as for computing eigenvalues from principal components analyses (e.g., EIGENSTRAT [14]) and imputation of variants (e.g., IMPUTE2 [15]) are not listed.

2.2.1 PLINK

This is a suite of tools designed to conduct genome association analyses, including both gene-gene and gene-environment interactions [16]. The primary interaction analyses are based on logistic and linear regression. They can accommodate both gene-gene and gene-environment interactions on the genome-wide scale or on a smaller number of variants by creating subsets of genetic variants to test again each other or the environmental factor. The program has the flexibility to conduct both explicit tests for interaction by testing for the significance of the interaction term directly in the regression model or a non-explicit joint test by testing the main and interactions effects. There is also a faster option for conducting genome-wide gene-gene interactions (*fast-epistasis*) based on the Z-score for the differences in OR for SNP-SNP combinations between cases and controls or for cases alone (case-only test).

2.2.2 CASSI

This is a software package that is specifically designed to conduct genome-wide gene-gene interaction analyses in a computationally efficient manner ([17]; <https://www.staff.ncl.ac.uk/richard.howey/cassi/index.html>). This package corrects a minor error in the Wu et al. statistic [18] in the calculation of the variance for estimated rather than observed haplotypes and in the *fast-epistasis* variance originally implemented in PLINK.

2.2.3 BOOST

The Boolean Operation-based Screening and Testing (BOOST) program was designed to efficiently screen and then explicitly test for genome-wide gene-gene interactions [19]. The screening phase involves a non-iterative procedure to approximate the likelihood ratio and then all variant pairs that survive this screening are subjected to a classical likelihood ratio test in the testing phase.

2.2.4 MDR

The Multifactor Dimensionality Reduction (MDR) software package [20] is designed to conduct data mining on discrete variables and can be used to detect both gene-gene and gene-environment interactions and dichotomous outcomes [21]. The traditional MDR approach is a non-explicit test for interaction as it is a non-parametric test that combines factors that may be interacting in order to best discriminate the subjects among the dichotomous outcome. An extension of the MDR has been developed that incorporates a permutation-based approach that can explicitly test for interactions [22]. A recent extension to the MDR has

implemented a t -test approach that allows for quantitative outcomes [23]. The MDR method, however, is designed primarily for smaller sets of markers, but parallel computing could be utilized to conduct a genome-wide analysis.

3 Methods

3.1 Quality Control (QC)

No analysis can be successful without high-quality data. The specific steps of the genotype QA/QC will depend on the type of assay used to generate the genotypes. These range from single-variant assays based on PCR, whole genome microarray genotyping, and whole-exome and whole-genome sequencing. A brief outline of the QC steps for each is outlined below.

3.1.1 Single-Variant QC

The primary steps are to assess the overall performance of the individual genotyping assays through examination of the variant call rate (variant call rate = total number of genotype calls/total number of individuals genotyped) and Hardy-Weinberg Equilibrium (HWE). While the thresholds chosen to eliminate variants can be arbitrary, typically one would look for variant call rates >98% (which should be examined in cases and controls separately if conducting a case-control analysis to ensure no bias due to differences in call rates between cases and controls) and HWE p -values $>10^{-4}$ which if conducting a case-control study are assessed only in controls. Without the benefit of genome-wide genotype data it is impossible to assess the data for population stratification, but adjustments can be made in the analysis (if using regression-based methods) for relevant covariates that may capture potential stratification such as self-reported race/ethnicity.

3.1.2 Microarray QC

As with single-variant QC, one will examine both the individual variant call rates and HWE to ensure that each variant probe is generating high-quality genotype data with similar thresholds applied as mentioned above. However, additional steps can and should be taken into account. The individual subject call rates should be examined first to determine if there were general problems with the individual array and/or DNA. These call rate thresholds may range from 93% to 98% and are often suggested by the array manufacturer-based past performance of the array (subject call rate = total number of genotype calls for an individual subject/total number of variants attempted to be genotyped). Poorly performing subjects should be removed prior to any downstream QC steps. Population stratification should be assessed using a genome-wide procedure such as principal components analysis to determine if there are slight variations in the genotype frequencies between subpopulations within your dataset. This procedure can detect any systematic differences that may be due to differences in

allele frequencies arising ancestry differences, but also due to exper- 269
imental/processing differences (e.g., plate effects). If significant 270
principal components (PCs) are detected it is suggested that these 271
PCs be adjusted for in the analysis. If the analysis assumes unrelated 272
subjects it is suggested that the dataset be examined for cryptic 273
relatedness using a procedure such as estimate pairwise identity-by- 274
descent (IBD). This pairwise measure that is often used is \hat{p}_i . Again, 275
the threshold for identifying cryptically related subjects is arbitrary 276
one often chooses a threshold ranging from 0.125 to 0.2 and then 277
eliminates one of the two subjects in this cryptically related pair. 278
This can be done randomly, or one may want to eliminate the 279
subjects based on the subject call rate (eliminating the subject 280
with a fewer genotype calls) or if it is a case-control study and a 281
cryptically related pair is comprised of a case subject and a control 282
subject it may be beneficial to eliminate the control if cases are in 283
short supply. 284

To assess whether or not the QC steps that have been taken are 285
successful prior to conducting an interaction analysis, it would be 286
beneficial to conduct a genome-wide single-variant analysis and 287
examine QQ plots and/or estimate λ from the data after 288
adjusting for PCs and other covariates. QQ plots can be generated 289
using an R script such as qqman.r ([https://CRAN.R-project.org/](https://CRAN.R-project.org/package=qqman) 290
[package=qqman](https://CRAN.R-project.org/package=qqman)) and λ estimated using PLINK. Deviations from 291
the expected line on the QQ plot are not expected except at the tail 292
(i.e., the true positives) and with deviations along much of the 293
expected line being an indication of residual population stratifica- 294
tion. λ estimates greater than 1.05 are routinely seen as indicators of 295
population stratification and additional PCs should be adjusted for 296
until the lambda value falls below this threshold. This assessment is 297
typically done in single-variant analyses prior to any interaction 298
analyses. 299

3.1.3 Sequencing QC

Specific workflows for alignment, variant calling, and variant 301
QA/QC and filtering are described in detail elsewhere (*see* Ref. 302
[24](#) for a detailed step-by-step pipeline covering the major sequenc- 303
ing analysis tools). Briefly, a standard analysis pipeline would start 304
by aligning the FASTQ raw sequence reads to a reference genome 305
using the Burrows-Wheeler Aligner (BWA, [[25](#)]). Then converted 306
to BAM format, sorted, indexed, PCR duplicates marked and then 307
merged into one BAM file using SAMtools [[26](#)]. Finally, the align- 308 [AU3](#)
ments in the BAM file can be locally realigned around insertion/ 309
deletions, recalibrated and variants called using HaplotypeCaller in 310
the Genome Analysis Toolkit (GATK, [[27](#), [28](#)]). 311

Variant QC can utilize a variety of different metrics, but an 312
example of one approach is how we conducted our QC in our 313
whole-exome sequence analysis of a family segregating asthma [[29](#)]. 314
Variants were flagged (and not considered further in our analysis) if 315

they met any of the following criteria: three or more variants detected within 10 bp; four or more alignments map to different locations equally well; coverage of less than five reads; quality score <50; low quality for a particular sequence depth (variant confidence/unfiltered depth <1.5); and strand bias (Phred-scaled p -values using Fisher's Exact Test >200). There are many variations that can be employed in your QC pipeline, but the most important aspect is achieving the highest quality set of variants to retain in your analysis.

3.1.4 Linkage Disequilibrium Pruning

Reducing the number of variants considered in the analysis, regardless of genotyping method, can be accomplished by linkage disequilibrium (LD) pruning. LD is an indication of the correlation, or non-independence, between variants and can be measured using r^2 or D' . A typical r^2 threshold used to prune variants is 0.8, but a lower threshold can be used to eliminate more variants at the risk of excluding some that are independently informative. This procedure is particularly important when conducting a case-only gene-gene interaction analysis (discussed in Subheading 3.2.12) as the variants must not be correlated with each other.

3.2 Gene-Gene Interaction Analysis

The basic analysis is straightforward, the assessment of a deviation from an additive or multiplicative model containing two or more variants. This is traditionally assessed through regression-based modeling but many different methods are available and several different methods will be discussed below to exemplify this approach. One could divide these approaches into explicit vs. non-explicit tests for interactions, with explicit tests determining if the null hypothesis of the sum on the additive/linear scale or the product on the multiplicative scale of the joint effects of the two variants is contributing to the outcome with the alternative hypothesis being that the joint effect of the two variants is greater or less than the expected. Non-explicit tests are able to determine if the grouping of variants is associated with the outcome, but not necessarily that the variants together deviate from an additive or multiplicative model. Different inheritance models can be imposed on the genetic variants in the interaction models (*see Note 1*) and the issue of outcome scale is also an important consideration when interpreting interactions (*see Note 2*).

Regression-based approaches are attractive not only for their ability to explicitly detect interactions in the data, but because of the ability to adjust for multiple covariates. In a genome-wide setting this is important so that significant principal components can be adjusted to remove population stratification, but other potential confounders can be adjusted as well, including age, sex, etc. depending on the outcome of interest.

3.2.1 Logistic Regression

For dichotomous outcomes, a logistic regression model is one of the most straightforward approaches for testing an interaction, either explicitly or non-explicitly. In the following logistic regression model we are modeling the probability of our dichotomous outcome (p) using an intercept (β_0), two genetic variants (SNP1 and SNP2, for example) and their respective main effects (β_1 and β_2) and their interaction effect (β_3):

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \beta_1 \text{SNP1} + \beta_2 \text{SNP2} + \beta_3 \text{SNP1} \times \text{SNP2}$$

In this model, we can test for the significance of the interaction effect of the two variants by testing the null hypothesis that $\beta_3 = 0$. This is an explicit test for interaction in that we are testing for an interaction after adjusting for the main effects (i.e., independent effects) of the two genetic variants. This model will produce an effect estimate (β_3) which can easily be converted to an OR by $\exp\beta_3$. It should be noted that the main effects should not be interpreted as the main effect of the variants (or the variant and the environmental factor) since they are adjusted for the interaction term in the model.

Alternatively, we can test for the significance of the joint effects of the two variants by using a 2 degree of freedom likelihood ratio test of the full model against a model in which there is neither an interaction term nor main effect term for one of the SNPs:

$$\log\left(\frac{p}{1-p}\right) = \beta_0 + \beta_1 \text{SNP1}$$

While this model contains an interaction term this is a non-explicit test for interaction as we are testing the significance of the main effect of one of the SNPs and interaction term combined. This may be seen as a less desirable test but this can be more powerful for detecting significant signals [30]. However, this joint test can be significant if either of the main effects is strong enough or if both main effects are strong enough without a significant interaction term (β_3). As will be discussed later, one may want to screen combinations of variants for joint effects and then follow up for explicit tests of an interaction only on the limited set of variants demonstrating significant joint effects.

When reporting the results of regression-based analyses, it is important to report not only the significance of the interaction (p -value), but also the parameter estimate of the interaction as well as those of the main effects (*see Note 3*).

3.2.2 Linear Regression

For continuous outcomes, linear regression can be used to test for interactions in much the same way as logistic regression for dichotomous outcomes. β s are not converted to ORs, but are rather

interpreted directly with the sign of β_3 , an indication of the direction of the interaction, and the p -value for the likelihood ratio test of the full model versus one without the interaction term an indication of the significance of the interaction term. A two degree of freedom joint test can also be constructed.

3.2.3 Nonparametric Analyses

The most popular nonparametric method for interaction analysis are the multidimensional reduction methods. These are based on the idea that across a contingency table of two or more genetic variant genotype combinations or genotype and discrete environmental factor combinations, each square in the contingency table can be divided into increasing risk and decreasing risk in the context of the phenotype. These collapsed contingency tables can then be tested for association with the phenotype. This is typically done by splitting the data into a training set (to build the collapsed contingency table) and a testing set to determine the classification error. The data are split multiple times and the model is assessed by the classification error and prediction error estimated from these multiple cross-validations [21]. Due to the fact that this is a nonparametric method, a p -value is obtained for the set of factors included in the model, but an effect estimate is not. The significance of the model is interpreted as the significance of the group of factors (genetic variants or genetic variant(s) and environmental factors) but cannot be interpreted as direct evidence of an interaction.

The model-based MDR (MB-MDR) methodology is an extension of the classic MDR framework that allows for the direct testing of interactions through the use of the Wald statistic on high and low-risk genotype categories [22, 31]. The significance of the interaction is then tested using permutation testing of the maximum Wald statistic and can be considered an explicit test of an interaction.

3.2.4 Multiple Test Correction

One of the biggest challenges in conducting genome-wide analysis in general, and more specifically interaction analysis, particularly gene-gene interactions, is the multiple testing problem. Given that hundreds of thousands or millions of variants are considered in most genome-wide studies, the chance of detecting a false positive is immense. This is even more apparent when multiple environmental factors are considered and most problematic when all pairwise variant interactions are considered in an exhaustive gene-gene interaction search. Traditionally, genome-wide studies have adopted a Bonferroni correction to adjust for the number of variants analyzed, but this can be overly conservative, especially for dense maps of markers in high LD. However, other procedures such as the False Discovery Rate (FDR; [32]), LD-based variant counting to account only for the number of independent tests [33] and permutation testing [34, 35] have been adopted. Permutation

testing, while attractive because it accounts for the number of independent tests by maintaining the LD structure in the permuted data, is computationally intensive, even more so when all pairwise variant combinations need to be considered in thousands of permuted datasets. Therefore, the FDR or LD-based methods are more attractive alternatives.

3.2.5 Genome-Wide (Exhaustive) Approaches

A typical genome-wide association study will genotype between 100,000 and 1,000,000 single-nucleotide polymorphisms. While exponential advances in computational power mean that the computing power to run the analysis on the pairwise tests required for a panel of 500,000 SNPs (2.5×10^{11} tests) is not insurmountable and with parallel computing relatively quick, the major hurdle is power. While large multi-center consortium studies mean that the number of subjects for many subjects has greatly increased, with studies reaching 50,000 or more subjects depending on the frequency of the disease or phenotype, these may still not be powerful enough to detect significant interactions on the genome-wide scale. In a study of genome-wide data obtained from the Resource for Genetic Epidemiology Research on Adult Health and Aging (GERA), using 45,171 subjects for ten phenotypes and conducting an exhaustive search for interactions, we failed to identify any genome-wide significant interactions, suggesting that we were underpowered to detect interactions with apparently weak effect sizes [36].

As demonstrated in Table 1, for a genome-wide gene-gene interaction study of 200,000 markers using an exhaustive approach, requires more than 30,000 cases and 30,000 controls to achieve genome-wide significance for two loci each with an MAF of 0.2 and an interaction OR of 1.2. All power calculations presented in this chapter were performed in Quanto [37]. The power calculations assumed a population risk of the disease of 0.1 and a log-additive mode of inheritance and a main effect of each variant of 1.2. For a case-control study with 5000 cases and 5000 controls, a reasonably sized case-control GWAS, and 200,000 markers genotyped, the minimal detectable OR is just over 1.6 (Table 2). For an interaction OR of 1.5, the number of markers considered would need to be reduced to around 2000 to detect a significant interaction among all pairs of interactions (4,000,000) and at an interaction OR of 1.2, only one interaction can be considered as only a nominally associated interaction is detectable (Table 3).

This lack of power can be daunting, but as detailed in the following sections there are several ways to increase power by reducing the number of tests through filtering or using biologically informed combinations of variants.

t.1 **Table 1**
Power of interaction analysis

t.2	Number of cases and controls						
t.3	Interaction type	5000	10,000	20,000	30,000	40,000	50,000
t.4	G-G	0.000	0.003	0.164	0.654	0.940	0.995
t.5	G-E	0.024	0.032	0.317	0.717	0.928	0.988

t.1 **Table 2**
Power of case-control and case-only gene-gene interaction analyses

t.2	intOR	Case-control	Case-only
t.3	1.1	0.000	0.000
t.4	1.2	0.000	0.013
t.5	1.3	0.004	0.547
t.6	1.4	0.076	0.992
t.7	1.5	0.388	1.000
t.8	1.6	0.782	1.000
t.9	1.7	0.961	1.000
t.10	1.8	0.996	1.000
t.11	1.9	1.000	1.000
t.12	2.0	1.000	1.000

3.2.6 Data Reduction Approaches

In order to overcome the multiple testing burden of a genome-wide screen for all gene-gene interactions, it may be advantageous to focus on a subset of SNPs that may have a one or more properties that may make the interaction analysis more likely to detect a significant interaction. If we reduce our set of 200,000 SNPs down to 2000 we can now detect an interaction OR of 1.5 using a set of 5000 cases and 5000 controls (Table 3). Below I outline a series of methods to reduce the set of SNPs considered.

Filtering by Allele Frequency

The simplest method to increase power to detect interactions is to impose a filter by minor allele frequency (MAF). Depending on the sample size and subsequent power estimates, it may be advantageous to filter out all SNPs with MAFs less than the power to detect a reasonable interaction effect size (e.g., all SNPs with an MAF <0.2).

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Table 3
Power of gene-gene interaction analysis by interaction OR and number of markers

# Markers	intOR = 1.2	intOR = 1.4	intOR = 1.5
200,000	0.000	0.076	0.388
100,000	0.000	0.107	0.465
20,000	0.000	0.221	0.649
10,000	0.001	0.289	0.724
2000	0.004	0.488	0.869
1000	0.008	0.584	0.914
200	0.037	0.793	0.975
100	0.066	0.864	0.988
20	0.219	0.966	0.999
10	0.337	0.986	1.000
5	0.488	0.995	1.000
2	0.713	0.999	1.000
1	0.865	1.000	1.000

Filtering by Marginal
Effects

While a significant marginal and/or main effect is not required to detect a significant interaction, there are few reports of significant interactions without also having detectable main effects. Therefore, it may be advantageous to filter based on marginal effects (i.e., single-variant effect) and only include those variants that have, for example, a single-variant association p -value of <0.05 (nominal significance threshold). This should reduce the number of SNPs to approximately 5% of the starting number (for our 500,000 SNPs this would result in 6.25×10^8 tests). This approach may be too stringent, as it requires both variants be nominally significant. An alternative approach would be to select a set of SNPs reaching a predefined significance threshold (e.g., $p < 10^{-4}$) and testing this set of SNPs against all other SNPs. In our example with 500,000 SNPs we would select ~50 SNPs with $p < 10^{-4}$ and test for an interaction with the remaining 499,450 for a total of 2.5×10^{-7} SNPs. This balances the requirement that there be some marginal effect of one of the SNPs with being able to detect interactions with SNPs showing no marginal effects but having a significance effect on disease only in the presence of a second SNP.

Candidate Gene
Approaches

Filtering based on prior evidence of a gene's involvement in a particular disease is another approach to reducing the search space for interactions. Candidate genes could be selected by systematically

reviewing the literature or for some diseases databases of candidate genes exist based on association studies, linkage analyses, and/or expression studies. For example, there are databases for preterm birth [38], preeclampsia [39], and non-syndromic hearing loss [40] for which one can obtain lists of candidate genes. By annotating variants to their respective genes using programs such as Annotvar [12] only those SNPs annotated to the set of candidate genes can be selected for inclusion in the interaction analysis. As for the marginal effects, it may be advantageous to consider those SNPs within candidate genes against all other SNPs genotyped. In this way, it is possible to detect interactions with SNPs in novel genes not previously identified to be associated with the disease of interest.

Filtering by Function

Filtering by the effects of specific variants is another approach to reducing the search space. Again, variants can be annotated, but now the selection may be made based on being a coding variant or a splicing variant that is more likely to be functional. Other biological information could also be utilized such as examining interactions between SNPs in genes known to be involved in protein-protein interactions with the rationale being that SNPs in these biologically interacting genes are more likely to also show evidence of a statistical interaction. Similarly, one could examine interactions between variants within transcription factors and those within their binding sites.

3.2.7 Multistage Approaches

In order to maintain a genome-wide approach but overcome the hurdle of the immense multiple testing problem, it may be beneficial to employ a multistage approach. This is possible in situations in which a study has a large sample size, but still not adequately powered to detect significant interactions genome-wide. In a study of asthma, we first screened all pairwise interactions (9.1×10^{10}) in a small subset of the data and then carried through all interaction with a suggestive significance ($p < 10^{-5}$) to a follow-up stage of independent subjects and then attempted to replicate the top SNPs in a third set of independent subjects [41]. While this approach did not identify any genome-wide significant interactions, it did identify a candidate interaction between SNPs in two regions of the genome. The major advantage of this type of approach is that it allows for an unbiased examination of the SNPs without relying on previously reported biological and/or association data.

3.2.8 Gene-Based Interaction Tests

Another approach to reduce the multiple testing burden and combine information across multiple variants is to conduct a gene-based test of interaction. By considering each gene as a unit in the interaction rather than each individual variant, the total number of interactions considered is significantly reduced, thus increasing power. One such approach combines interaction p -values across all

combinations of genetic variants in two genes into a single 581
gene-gene interaction p -value that also accounts for linkage 582
disequilibrium [42]. 583

It should be appreciated that these gene-based tests can be 584
applied to the setting of rare variant which are increasingly being 585
studied. Extensions to rare variant tests to incorporate gene-gene 586
and gene-environment interactions have been developed and 587
include SKAT [43] which can handle gene-gene interactions and 588
iSKAT [44] and rareGE [45] for gene-environment interactions. 589

3.2.9 Replication

Regardless of the analysis approach taken to identify gene-gene 591
interactions, the gold standard is to conduct a replication analysis 592
using an independent dataset. While fairly common for single- 593
variant association studies, this is less routinely followed for gene- 594
gene interaction analyses, as demonstrated in our systematic review 595
of asthma gene-gene interactions where only 15.2% of interactions 596
were attempted to be replicated [46]. The challenge is often iden- 597
tifying an appropriate replication dataset in which both variants 598
were genotyped. Through the use of imputation the variant to be 599
replicated could be imputed if they were not directly in the inde- 600
pendent dataset. 601

A challenge for any replication of a genetic effect is the direc- 602
tionality of the effect. Differences due to the populations selected 603
that can alter the minor allele frequencies and linkage disequilib- 604
rium structure can result in differences in both the magnitude and 605
direction of effect [47]. This is amplified when looking at two or 606
more loci as the probability of subtle differences can result in 607
differences in the direction of effect when looking at interactions. 608

3.2.10 Meta-Analysis

This data analysis technique is commonly used to pool data across 610
multiple studies and increase evidence for an association with a 611
genetic variant. Meta-analysis is an attractive analytical technique 612
because it can be used to increase power by substantially increasing 613
the total sample size. While relatively straightforward for single- 614
variant analyses using either fixed or random effects models [48] or 615
by combining p -values [49], this is not always the case for gene- 616
gene and gene-environment interactions. Differences in the analyt- 617
ical strategy and the way in which the results are presented may 618
make meta-analyses more challenging. The p -value approach does 619
not take into account the effect estimates, but this may be a better 620
first approach as it can be used on a much broader set of analytical 621
strategies such as MDR, regression-based approaches, and Random 622
Forrest methods, for example. Meta-analysis methods that account 623
for the effect estimate are more attractive because they can account 624
for heterogeneity in both the effect estimates and between 625
populations. 626

627

3.2.11 *Case-Only Approach*

The case-only study design to detect interactions was first described for gene-environment interactions [50, 51]. This design is premised on the idea that the variant is independent of the environmental exposure in the population. In the presence of a gene-environment interaction, there would be an association between the variant and the environmental exposure among cases only and this can be most easily tested using the regression-based approaches described above (linear or logistic depending on the environmental exposure being investigated); however, it should be noted that when using the case-only design the main effects of the variant and the environmental exposure cannot be determined as this is only possible using a case-control design.

The case-only design is more powerful than the case-control design (Table 4; 5000 cases and 5000 controls for case-control and 5000 cases for case-only, main effects ORs of 1.2 for both the environmental factor and genetic variant, a frequency of the environmental exposure of 0.25 in the population and MAF of 0.2, population risk of disease of 0.1 and 200,000 genetic markers). The major caveat is that there is an assumption that there is independence between the genetic variant and the environmental factor, i.e., that there is no association between these factors among the source population. Violation of this assumption can lead to biased estimates of the OR and there is a recommendation that a case-only gene-environmental interaction analysis only proceed for environmental factors with population-specific data [52]. This independence may be difficult to establish in population-based data, so independence could be tested for among controls (e.g., when a

t.1 **Table 4**
t.2 **Power of case-control and case-only gene-environment interaction**
t.3 **analyses**

t.2	intOR	Case-control	Case-only
t.3	1.1	0.000	0.001
t.4	1.2	0.002	0.063
t.5	1.3	0.038	0.564
t.6	1.4	0.211	0.962
t.7	1.5	0.541	1.000
t.8	1.6	0.829	1.000
t.9	1.7	0.959	1.000
t.10	1.8	0.994	1.000
t.11	1.9	0.999	1.000
t.12	2.0	1.000	1.000

case-only analysis is being conducted among a case-control study) 655
 or replication of the finding among an independent case-control 656
 study. This is of particular importance when the environmental 657
 factor could be under genetic influence such as BMI. The substan- 658
 tial power gain of the case-only study makes this analysis strategy 659
 relatively attractive, but this is somewhat diminished by the strong 660
 independence assumption and the need to test this assumption in a 661
 set of population-based samples. 662

The case-only approach can also be utilized to detect gene-gene 663
 interactions [53]. In the presence of an interaction between two 664
 genetic variants there will be an association between these genetic 665
 variants among cases. This is most easily tested using a logistic 666
 regression model using one of the genetic markers as the dependent 667
 variable. As with case-only gene-environment interaction analyses, 668
 only the interaction can be examined and not the main effects of the 669
 genetic variants. Independence of the genetic variants in the source 670
 population is assumed, but this is much easier to achieve given the 671
 large amounts of genome-wide genotype data available for many 672
 diverse populations (e.g., HapMap [54]) or the use of LD pruning 673
 among controls, if available (*see* Subheading 3.1.4). 674

3.2.12 Three or More Variants

Gene-gene interactions are not limited to pairwise interactions. 676
 Higher order interactions involving three or more variants have 677
 been reported (e.g., renin-angiotensin system SNPs and hyperten- 678
 sion [55]) and can be modeled in the regression framework. The 679
 problem is that as higher order interactions are considered the data 680
 become more and more sparse and the power to detect a significant 681
 interaction decreases. One approach to detecting higher order 682
 interactions would be to use a nonparametric, non-explicit method 683
 such as MDR to screen for potential higher order combinations of 684
 variants and then subsequently explicitly test for the interaction on 685
 a subset of the best performing combinations using a regression- 686
 based method. 687

3.3 Gene- Environment Interaction Analysis

At the core, the analysis of gene-environment interactions does not 689
 fundamentally differ from gene-gene interactions. The analysis can 690
 be conducted using the regression-based or nonparametric 691
 approaches described above with the same caveats for testing for 692
 explicit vs. non-explicit interactions in the dataset. Joint, or non- 693
 explicit, tests may be used to screen the data to detect potential 694
 interactions, followed by explicit tests to determine if an interaction 695
 exists between the identified variant and the environmental factor. 696
 In the logistic regression framework, one of the genetic variant 697
 variables is replaced by the environmental exposure of interest and 698
 the significance of the interaction β is tested in the model. Using the 699
 non-explicit MDR approach, the environmental variable is entered 700
 into the algorithm with the caveat that the environmental factor 701
 must be categorical in order for the reduction algorithm to work. 702

The power to detect gene-environment interactions suffers from the same lack of power on the genome-wide scale as gene-gene interactions. More than 30,000 cases and controls are required to detect an interaction OR of 1.2 (Table 1) under similar parameters as the gene-gene interactions analysis (main effects of OR = 1.2 for gene and environmental factors, MAF of 0.2, environmental prevalence of 0.25, population risk of disease of 0.1 and 200,000 genetic markers tested). Using 5000 cases and 5000 controls, the smallest detectable OR is just under 1.6 (Table 4). The same approaches to increase power by reducing the search space can be applied to the search for gene-environment interaction analyses, through all of the data reduction techniques described for gene-gene interactions.

It cannot be stressed enough that replication is the key to describing true positive gene-environment interactions, as was discussed for gene-gene interactions. As mentioned previously for gene-gene interaction replication, if the exact genetic variants are not directly genotyped in the independent dataset, imputation can be used to estimate the genotypes to be replicated. While this will work for the genetic variants in gene-environment interactions to be replicated, this is not the case if there is not comparable environmental data in the independent dataset. This can make the identification and selection of an appropriate dataset for replication more challenging for gene-environment interactions, but makes the replication of the findings no less important.

3.3.1 Modifiable Environmental Factors

One attractive aspect of identifying and describing gene-environment interactions for a complex disease is that this gives us the possibility of potentially modifying one contributing factor to a disease. At this point in time, inherited genetic variants are not modifiable, but if, for example, a significant gene-environment interaction were identified for obesity with a genetic variant and high saturated fat diet, individuals carrying the risk variant could be more strongly encouraged to reduce their saturated fat intake. Despite the general benefit we could all gain from reducing our saturated fat intake, this may be more effective if it were targeted to individuals based on their genetic profile and increased risk for obesity when both factors are present (beyond the additive main effects).

3.3.2 Gene-Gene- Environment Interactions

We should not think of gene-gene and gene-environment interactions as being mutually exclusive. As with higher order gene-gene interactions, interactions involving multiple genetic variants and an environmental factor are possible to model. The same issues with power due to sparsity of data and the number of unique combinations of factors apply, but given sufficient sample size these types of interaction models can be tested. It may be more important in this setting to attempt a screening step using a non-explicit approach and then apply an explicit test only to the set of interactions that surpass an initial significance threshold.

3.4 Conclusions

I have outlined several methods for conducting interaction analysis to detect both gene-gene and gene-environment interactions. However, it should be clear that there is no optimal method to detect either type of interaction. The method(s) chosen are often dictated based on the type of data you have available (e.g., case-control, case-only), the number of markers you have genotyped, and the number of subjects you have included. I strongly recommend reporting the results of explicit tests for interaction as this will greatly improve ability of other groups to attempt to replicate your results and meta-analyze where appropriate, but the use of non-explicit tests for interaction can be extremely useful to initially screen large numbers of interactions and when sample sizes are limited. We must continue to invest time and resources into identifying interactions in genome-wide data as this will help us to fill in the missing heritability gap and better understand the genetic architecture of complex traits.

4 Notes

1. Inheritance models: It should be noted that as with single-variant approaches, inheritance models can be imposed on the variants that include additive, multiplicative, dominant, recessive, and overdominant. These inheritance models can be imposed on the variants in the interaction model independently and can be considered in a combinatorial fashion for variant 1 and variant 2 (e.g., additive \times additive, additive \times multiplicative, additive \times dominant, etc.). However, if this is done, care must be taken to account for this additional multiple testing. These models can be problematic when they are misspecified as they can reduce power which makes nonparametric approaches attractive.
2. Scale: The scale on which the outcome is measured or evaluated can influence whether or not an interaction exists between two variants or between a variant and an environmental factor. It needs to be kept in mind that there may be a monotone transformation of the outcome that could remove the interaction. For example, on the odds ratio scale, an interaction may exist between two SNPs (coded dichotomously as in a dominant model) in which the OR for having dominant alleles at both SNPs is greater than the sum of the ORs for having one dominant allele at each SNP. However, on the log(OR) scale this interaction is removed and termed a removable interaction. If there is no monotone transformation that can remove this interaction it is termed essential. The method to detect these types of interactions is described in a paper by Wu et al. [56].

3. Reporting of interaction results: As was previously outlined in our paper [46], there are several recommendations for how results of interaction analyses are reported in order to increase the interpretability and replicability of the interaction. Effect estimates should be provided so that both the strength and direction of the interaction can be assessed. If a regression-based approach is used, parameter estimates of the main effects and the interaction term should be provided. If a nonparametric approach is used, such as MDR, effect estimates are not produced, however, counts of cases and controls for the contingency table of genotype combinations should be provided. This will allow for a better assessment of the interaction and the possibility to incorporate the data into a meta-analysis.

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