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Corresponding Author	Family Name	Michailidou
	Particle	
	Given Name	Kyriaki
	Suffix	
	Division	Department of Electron Microscopy/Molecular Pathology
	Organization	Cyprus and Centre for Cancer Genetic Epidemiology, The Cyprus Institute of Neurology and Genetics
	Address	Nicosia, Cyprus
	Division	Department of Public Health and Primary Care
	Organization	University of Cambridge
	Address	Cambridge, UK
	Email	kyriakimi@cing.ac.cy
Abstract	Meta-analysis is a statistical technique that is widely used for improving the power to detect associations, by synthesizing data from independent studies, and is extensively used in the genomic analyses of complex traits. Estimates from different studies are combined and the results effectively provide the power of a much larger study. Meta-analysis also has the potential of discovering heterogeneity in the effects among the different studies. This chapter provides an overview of the methods used for meta- analysis of common and rare single variants and also for gene/region- based analyses; common variants are mainly identified via genome-wide association studies (GWAS) and rare variants through various types of sequencing experiments.	
Keywords (separated by '-')	Meta-analysis - Common variants - Rare variants - Aggregation analysis - Single variant analysis - GWAS - NGS	

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# Chapter 6

## **Meta-Analysis of Common and Rare Variants**

### Kyriaki Michailidou

#### Abstract

Meta-analysis is a statistical technique that is widely used for improving the power to detect associations, by 5 synthesizing data from independent studies, and is extensively used in the genomic analyses of complex 6 traits. Estimates from different studies are combined and the results effectively provide the power of a much 7 larger study. Meta-analysis also has the potential of discovering heterogeneity in the effects among the 8 different studies. This chapter provides an overview of the methods used for meta-analysis of common and 9 rare single variants and also for gene/region-based analyses; common variants are mainly identified via 10 genome-wide association studies (GWAS) and rare variants through various types of sequencing 11 experiments.

Key words Meta-analysis, Common variants, Rare variants, Aggregation analysis, Single variant 13 analysis, GWAS, NGS

#### 1 Introduction

Meta-analysis provides a powerful approach to combine data from 16 different resources. It has been widely used in genomics in order to 17 increase the power of single studies to detect associations with a 18 specific trait or disease of interest [1, 2]. Sharing of genotype data 19 between studies is not always possible, even within well-established 20 collaborations, thus the need for alternative approaches for com-21 bining the effects from different studies has led to the wide use of 22 meta-analyses in the field. The different meta-analysis techniques 23 use summary statistics, are easy and quick to perform, are powerful, 24 and practically provide the power of a much larger study. 25

Meta-analysis has been extensively used in order to assess 26 Genome-wide association study (GWAS) data for millions of geno-27 typed or imputed SNPs [2]. GWAS provide a cost-effective method 28 for assessing the effect of common genetic variation across the 29 genome [3, 4] and have been widely informed using imputation 30 to publicly available genotype reference panels such as the HapMap 31 [5], the 1000 Genomes Project Consortium [6], and the 32

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Haplotype Reference Consortium [7]. Over the last decade 33 thousands of common variants, which are associated with complex 34 diseases or traits, have been identified through imputation and 35 meta-analyses of GWAS [8, 9]. More recently, custom arrays that 36 focus on the replication of rare variants from sequencing experi-37 ments have been developed [10]. The identification of rare variants 38 through genotyping arrays and imputation has been more prob-39 lematic compared to common variants [11]. With the tremendous 40 advances in next generation sequencing (NGS) technologies [12], 41 it is now feasible to conduct large-scale whole-genome, whole-42 exome, and targeted sequencing experiments. Recently, novel sta-43 tistical techniques have been developed for the analysis of the rare 44 variants as single entities and also for their collective analysis in gene 45 or regional tests. As in the common variant analysis, meta-analysis 46 techniques will aid the increase of power to detect associations with 47 rare variants. In the following sections, the different aspects of 48 meta-analysis for common and rare variants are discussed. 49

Meta-analysis usually begins with the individual studies sharing Meta-Analysis 50 summary statistics for each variant, including a regression estimate, Pre-Steps and Quality 51 standard error, p-value, sample size, imputation accuracy, and Control 52 minor allele frequency (MAF). There are various steps that need 53 to be taken into account before performing the meta-analysis in 54 order to minimize bias. Studies need to harmonize their quality 55 control measures and perform the analysis in the same/comparable 56 way before being able to combine the results. Standardized quality 57 control measures that need to be followed in each of the participat-58 ing studies, for example in GWAS, include the removal of low call 59 rate individuals and variants, removal of variants with genotype 60 frequencies deviating from those expected under Hardy-Weinberg 61 Equilibrium and variants with poor clusterplots [4]. Parameters for 62 adjustments need to be set upfront so that the analyses are per-63 formed in a comparable way. Principal components should be 64 calculated and adjusted for in the analyses [13] and appropriate 65 genomic control can be applied to individual studies [14] in order 66 to minimize bias due to population stratification. If the studies 67 perform imputation they need to use the same reference panel 68 and the same filters afterward (MAF and imputation accuracy) 69 [15]. Different programs produce different quality metrics that 70 need to be accounted for when performing the meta-analysis, for 71 example "info score" from IMPUTE2 [16], "R<sup>2</sup>" from MACH 72 [17] and BEAGLE [18]. The individual study effect estimates need 73 to be aligned to the same strand (usually more difficult for the 74 ambiguous SNPs with A/T and C/G genotypes) and variants 75 with large differences from the mean MAF need to be checked in 76 more detail. Short insertion/deletions (INDEL) that are now 77 being successfully imputed, using the 1000 Genomes Project data 78 as reference, can have different annotations across the different 79 imputation software that need to be carefully matched. For example 80 some software keep the actual alleles from the reference panel 81 whereas others provide the INDELs as D/I. Furthermore, in the 82 1000 Genomes Project reference dataset one base is subtracted 83 from the genomic location of the INDELs and this results in 84 different genomic locations for the same variants across the differ-85 ent reference resources. 86

If the data come from sequencing experiments they need to be 87 aligned on the same reference genome and low-quality variants 88 need to be removed before the meta-analysis is performed, to 89 avoid spurious associations due to sequencing errors. Furthermore, 90 special care needs to be taken when combining the results of 91 sequencing experiments that have been produced using different 92 technologies. Different depth/coverage of the regions of interest 93 or genotyping bias due to the differences in sequencing technolo- 94 gies can lead to the wrong conclusions. If the analyses are based on 95 genes or regions with variable thresholds the classifications for 96 SNPs/variants to be included in each gene/region need to be the 97 same (for example minor allele frequency threshold). Before 98 performing meta-analysis, a common statistical analysis plan 99 needs to be adopted to ensure compatibility of the results and to 100 aid a smooth execution [2]. 101

#### 2 Common Variants

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Different analytical approaches have been proposed and have been 104 extensively used for common variant meta-analyses [1, 2].

#### 2.1 Fixed Effects

The most widely used technique for meta-analysis of common 106 variants is the fixed-effects meta-analysis [1, 2]. The assumption 107 behind the fixed-effect meta-analysis is that there is a single common underlying genetic effect in the different studies. This has 109 been proven a powerful approach for discovery of common genetic 110 variants and provides practically the same information as a pooled 111 analysis of the raw data [19]. Different weights have been proposed 112 with the most optimal weight being the inverse variance [20]; other 113 weighting methods have also been used, such as the Mantel-114 Haesnzel method [21] or weights proportional to the sample 115 size. The weighted effect and variance are calculated as: 116

$$\overline{\beta} = \frac{\sum_{i}^{k} w_{i} \beta_{i}}{\sum_{i}^{k} w_{i}}$$
$$\overline{p} = \frac{1}{\sum_{i}^{k} w_{i}}$$

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where  $\beta_i$  is the effect estimate from each study,  $w_i$  is the weight assigned to each study and k is the number of studies. For the fixedeffects inverse variance meta-analysis,  $w_i = \frac{1}{v_i}$ , where  $v_i$  is the variance of each study [22]. The fixed-effect weighted test statistic follows a standard normal distribution:

$$\frac{\overline{\beta}}{\sqrt{\overline{\nu}}} \sim N(0,1)$$

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Random Effects When the underlying effects for each of the studies are assumed to 2.2 124 be different but drawn from the same distribution with variance  $\tau$ , 125 random-effects meta-analysis should be used. When there is no 126 presence of heterogeneity across the different studies, the random-127 and fixed-effects estimates will give approximately the same results. 128 The most common methods for calculating the variance of the 129 effect distribution are the method of moments [23] or likelihood-130 based methods [24]. Random-effects models have been used 131 mainly for the determination of the generalizability of the results 132 of the meta-analysis rather than for discovery purposes, as they are 133 less powerful [2]. Most commonly researchers report both fixed-134 and random-effects meta-analysis results. The classic random-135 effects meta-analysis follows the inverse-variance scheme with the 136 difference that the variance is now the sum of the within-study 137 variance  $(v_i)$  plus the between-studies variance  $(\tau^2)$  [22]: 138

 $w_i^* = \frac{1}{v_i^*}$   $v_i^* = v_i + \tau^2$   $\tau^2 = \frac{Q - (k - 1)}{\sum_i^k w_i - \frac{\sum_i^k w_i^2}{\sum_i^k w_i}}$   $Q = \sum_i^k w_i \beta_i^2 - \frac{\left(\sum_i^k w_i \beta_i\right)^2}{\sum_i^k w_i}$ 

where  $\beta_i$  is the effect estimate from each study,  $w_i$  is the weight assigned to each study, and k is the number of studies. The weighted effect and variance becomes 141

$$\overline{\beta}^* = \frac{\sum_i^k w_i^* \beta_i}{\sum_i^k w_i^*}$$

2.3 P Value and

Z Score Meta-Analyses

Meta-Analysis Methods

$$\bar{v}^* = \frac{1}{\sum_i^k w_i^*}$$

The test statistic then follows

$$\frac{\overline{\beta}^*}{\sqrt{\overline{\nu}^*}} ~ N(0,1)$$

Han and Eskin [25] observed that classical random effect mod- 143 els are underperforming even when there is heterogeneity present 144 among the different studies. They proposed an alternative method 145 for performing random effects meta-analysis whereby under the 146 null hypothesis it is assumed that there is no heterogeneity. Han 147 and Eskin [26] proposed the Binary Effects Assumption as another 148 method for random effects meta-analysis. This method is based on 149 two hypotheses; first that the effect is either present or absent in a 150 study and second that if the studies have an effect then the effect is 151 expected to be similar between the studies. A novel random effects 152 model, based on a kernel machine framework, has been proposed 153 by Shi et al. [27] for the meta-analysis of trans-ethnic studies. In the 154 presence of substantial heterogeneity between the results in the 155 different studies, further checking needs to be made to explore 156 the potential reasons behind this heterogeneity [25]. 157

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A more simplistic meta-analysis approach is the meta-analysis using 159 the *p*-values of the individual studies [28] and the test statistic takes 160 the form: 161

$$X_{2k}^2 = -2 \sum_{i}^{k} \log(p_i)^{\sim} X_{2k}^2$$

where k is the number of studies and  $p_i$  the individual study p-value. 162 Z score statistics-based meta-analysis [29] has also been used, the 163 test statistic can be derived using the p-values together with the 164 sample size information and direction of the effect. The Z-score can 165 be calculated using the following equation: 166

$$Z=rac{\sum_{i}^{k}Z_{i}w_{i}}{\sqrt{\sum_{i}^{k}w_{i}^{2}}}$$
 ~ $N(0,1),$ 

where

$$Z_i = \Phi^{-1} \left( 1 - \frac{p_i}{2} \right) \times \text{direction of effect}$$

where  $w_i$  is the square root of the sample size and  $\Phi$  is the standard 168 normal cumulative distribution function. Although these methods 169 are more straightforward to perform there is a substantial loss of 170 power as no information regarding the direction of the effects in 171

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each study is used (for the *p*-value meta-analysis), a combined effect 172 cannot be calculated and measures of heterogeneity cannot be 173 obtained [2]. 174

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2.4 Bayesian Meta-Although not as widely used as the other classical methods for 176 meta-analysis, different Bayesian methods have been also adapted 177 for different scenarios, with both fixed and random effects. Results 178 obtained from Bayesian meta-analyses are directly comparable 179 among/across different SNPs without the need for comparing 180 power and adjusting for differences in the frequencies among dif-181 ferent variants [30]. Choosing a suitable prior to be used for the 182 meta-analysis is an important issue and certain assumptions need to 183 be made. Results obtained via Bayesian meta-analysis include the 184 Bayes Factor (BF). The conventional cutoff for a test to be signifi-185 cant (BF > 10) is not sufficient for GWAS, not due to issues of 186 multiple testing as in frequentist testing but because the number of 187 truly associated variants we expect to have is small [30]. Bayesian 188 methods have also been developed for the meta-analysis of trans-189 ethnic GWAS data, where the studies are assigned into ethnic 190 clusters and the effects are assumed to be the same in each ethnic 191 group [31]. 192

#### 3 **Rare Variants**

The significant reduction in the cost of whole-exome and whole-195 genome sequencing in recent years has enabled large-scale sequenc-196 ing experiments to be conducted [7, 32]. Chip-based rare variant 197 experiments [10, 33] have also been performed for the assessment 198 of rare variation. 199

Single variant tests for rare variants are more challenging to 200 perform compared to common variants. The statistical tests need to 201 be adapted for rare variants as there is usually only a small number 202 of alternative allele counts and the current methods might not be as 203 accurate [11]. Score type statistics have been shown to be more 204 stable for rare variant association testing, especially for binary traits 205 [34]; Wald statistics can be too conservative and likelihood type 206 statistics too liberal. 207

Rare variant experiments are underpowered to identify single var-3.1 Aggregation 208 iants associated with modest effects, even within a large sample size Tests 209 [11]. Different methods have been proposed to increase the power 210 to detect associations by grouping variants into units of interest 211 (regions). A large number of different aggregation tests have been 212 developed to combine the effects of a subset of variants in order to 213 obtain a region-level test statistic [11]. Aggregation testing 214 increases the power to detect associations by combining the 215

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cumulative effects of rare variants and by reducing the number of 216 tests performed [11, 34]. 217

The various aggregation tests can be extended to include any 218 subset of variants of interest and for convenience, for the rest of this 219 chapter, these subsets will be referred to as a "region" of interest. 220 For example, in exome or whole-exome studies the region can be a 221 specific exon of a single gene, all the exons of a single gene or all the 222 exons of all the expressed genes in a genome. In the case of whole- 223 genome studies the region can also be a genomic region (for 224 example a sliding window of a pre-specified genomic length) or a 225 subset of variants belonging to the same category (for example 226 non-synonymous variants or variants that have a specific functional 227 annotation according to functional classification software [35]). 228 Some of the most widely used aggregation tests are discussed in 229 more detail elsewhere in this book (see Chapter 5). These include 230 the burden test or collapsing test, where a score is created for each 231 set of variants in a region, for each study sample, and then this is 232 compared to the disease/trait of interest. Collapsing tests can either 233 calculate a binary value (0 rare alleles or at least one rare allele) such 234 as CAST [36], or count the rare variants in each gene [37], or 235 calculate a weighted sum of the rare alleles in the region 236 [38]. Another form of collapsing test is the CMC test where the 237 variants are grouped according to their MAF and then CAST is 238 performed [39]. A different set of approaches are the variable 239 threshold (VT) methods, where the decision of the MAF threshold 240 for variants to be included in the region is obtained so that it gives 241 the most significant result [34, 40]. Other methods developed are 242 the variance-component (VC) tests that can detect associations in 243 regions when variants are allowed to have opposite effects such as 244 the C-alpha and SKAT [41, 42]. A combination of the burden and 245 variance-component tests has also been developed such as the 246 optimal unified SKAT-O, which takes the most significant linear 247 combination of the burden and SKAT tests [43]. The majority of 248 the different methods are applicable for quantitative measures, 249 binary traits, and survival analysis data. Even when aggregating 250 rare variants in regions of interest the power to detect associations 251 is still small [11] and thus appropriate techniques for the meta- 252 analysis of regional tests have been developed in order to increase 253 the power to detect associations. Both fixed- and random-effects 254 meta-analysis methods have been proposed for the regional tests. 255 As estimates based on individual rare-variant regression are not 256 stable the most optimal methods for meta-analysis of regional 257 tests have been based on the meta-analysis of score statistics. 258 259

#### 3.2 Meta-Analysis of Regional Tests

3.2.1 P Value and Z Score

Following a similar concept as the single variant meta-analysis the 260 most straightforward method for regional meta-analysis is a *p*-value 261 [28] or a *Z* score statistic [29] meta-analysis. A regional *p*-value or 262 *Z* score is obtained from each study and then these are meta-263 analyzed. This is an attractive method when effect estimates cannot 264

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be calculated or shared between the different studies. However, 265 these methods have been shown to suffer from substantial loss of 266 information, especially when no information of sample size and 267 direction of the effects is used, and thus these methods are not 268 widely used [44]. Another simple method has been proposed by 269 Lumley et al. [45] and is performed by the summation of the test 270 statistics for each individual study, though it is not as powerful if 271 variants are shared between studies. 272

More sophisticated methods that do not result in loss of infor-273 mation and provide results that are as powerful as a pooled analysis 274 of individual level data have also been proposed. In the following 275 sections a general framework for the meta-analysis of rare variant 276 regional association tests will be described. These tests are based on 277 the individual study score test summary statistics and the fact that 278 the regional test statistic can be reconstructed using the individual 279 variant score statistic from each study [44-48]. Fixed effects and 280 random effects meta-analysis methods have been proposed for the 281 most widely used regional tests. Individual studies share the sum-282 mary score statistic for each variant and an average variance-283 covariance matrix for the region [44-48]. Others have also pro-284 posed meta-analysis methods for aggregation tests that are recon-285 structed using the effects and standard errors from the usual 286 regression analyses together with the correlation matrix of the 287 individual variants (which can also be obtained from public 288 resources) [49]. This is an attractive method when score statistics 289 cannot be obtained but there is a substantial loss of information 290 since variants that do not produce valid effect estimates are not used 291 [49]. In the next sections score-based meta-analysis tests will be 292 described; these methods have been shown to be as powerful as the 293 analysis of the pooled genotype data [44, 46, 47]. 294

#### 3.2.2 Fixed Effects

The main assumption behind fixed effect meta-analysis for regional 296 tests is that there is a shared common genetic effect across the 297 different k studies. If we assume that there are j variants in the 298 region of interest, we can get the combined score  $(U_j)$  and com-299 bined variance  $(V_j)$  for each variant [44, 50]: 300

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$$U_j = \sum_i^k U_{(j,i)} \text{ and } V_j = \sum_i^k V_{(jj,i)}$$

$$w = (w_1, w_2, \dots, w_j)^T$$
301

where w is a vector of weights. If a variant is not present in a specific 302 study then the corresponding score and variance are set equal to 303 0. The regional score (U) and variance/covariance matrix (V) are 304 defined as 305

$$U = \left(U_1, U_2, \dots, U_j\right)^T$$

#### Meta-Analysis Methods

$$V = \operatorname{cov}(U)$$

The majority of the different methods use a normalized score 307 statistic (for quantitative traits) for each variant [34, 46, 47] 308 whereas Liu et al. [44] uses non-normalized scores. 309

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For the burden test, the assumption is that the combined score for 311 the region is the same across the different studies  $\beta_1 = \beta_2 = ... = \beta_k$ . 312 For testing the null hypothesis that  $\beta_k = 0$ , under the additive mode 313 of inheritance, the meta-analysis test statistic takes the form [46]: 314

$$Q_{M-Burden} = \frac{U^{*2}}{V^*} {}^\sim X_1^2$$

where  $U^* = w^T U$  and  $V^* = w^T V w$ .

Equivalently form Liu et al. [44]:

$$Q_{M-Burden} = \frac{w^T U}{\sqrt{(w^T V w)}} ~ N(O, 1)$$

Weights are usually based on MAF threshold cutoff (for example MAF < 0.05 or MAF < 0.01) or the Madsen-Browning [38] 318 weights, which up-weight rarer variants. 319

A VT test can be constructed by calculating the burden test statistic 321 at each MAF threshold (*p*): 322

 $Q_{M-VT} = \max_{p} Q_{M-\text{Burden}}(p)$ 

The *p*-value can then be calculated by comparing the test  $_{323}$  statistic to a multivariate normal distribution of U[34].  $_{324}$ 

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For the VC tests, the mean of the genetic variants in the region is 326 assumed to be the same across studies. The mean  $\mu$  of the variants in 327 the region is assumed to follow a multivariate normal distribution 328 with mean 0 and covariance matrix  $\tau_{w}W$ . For testing the null 329 hypothesis that the mean of the variant level effects  $\mu = 0$ , the 330 meta-analysis test statistic takes the form [44, 46]: 331

$$Q_{M-SKAT} = U^T W U$$

where *W* is an *jxj* diagonal matrix of rare variant-specific weights, 332 usually a function of the MAF, for example if *W* is a diagonal Beta 333 (MAF<sub>i</sub>,a<sub>1</sub>,a<sub>2</sub>) this is equal to the SKAT statistic [41] and if W is an 334 identity matrix then it produces the meta-analysis statistic for the 335 C-alpha test [42]. The p-value of the test statistic can then be 336 obtained by comparing the test statistic to a mixture of  $\chi_1^2$  distribu-337 tions,  $\sum_i^j \lambda_i \chi_{1,i}^2$  where  $\lambda_i$  is the *i*<sup>th</sup> eigenvalue of  $V^{1/2}WV^{1/2}$  338 [44, 46, 47].

Burden Tests

Variable Threshold (VT) Tests

Variance Components (VC) Test, SKAT, and SKAT-0

Burden Test

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An optimal unified test, SKAT-O [43], has been proposed; for this test the most optimal linear combination of SKAT and burden test is selected. The meta-analysis formula for SKAT-O takes the form [34, 43] 343

$$Q_{\text{hom-SKAT-O}} = \rho Q_{M-\text{Burden}} + (1-\rho) Q_{M-\text{SKAT}}$$

Intuitively if  $\rho = 1$  the test correspond to the meta-analysis of 344 burden test and if  $\rho = 0$ , to the meta-analysis of the SKAT statistic.  $\rho$  345 is calculated so that it produces the most significant result, the 346 p-value of the test can be obtained using a one-dimensional numerical integration [43]. 348

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3.2.3 Random-Effects Random-effects meta-analysis for regional tests assumes that the 350 genetic effects of the different studies are not the same but are 351 derived from the same distribution [51]. Heterogeneity between 352 and across studies is expected to be a bigger issue for rare variants 353 compared to common variants, as they are population specific 354 [51]. Random-effects models for aggregation tests have been 355 based on the Han and Eskin [25] single variant meta-analysis 356 method that has been shown to be more powerful compared to 357 other classical methods. 358

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Under the random effects model for the burden test, the combined 360 effects of the different studies are drawn from: 361

$$\beta_i = \mu + \xi_i, \quad i = 1, \dots, k$$

where  $\mu$  represents the average combined genetic effect among 362 studies and  $\xi_i$  represents the deviation of the effect of study k 363 from the mean  $\mu$ , and is assumed to follow a multivariate normal 364 distribution with mean 0 and variance  $\sigma$ . The test for  $\mu = 0$  and 365  $\sigma = 0$  takes the form [51]: 366

$$Q_{RM-Burden} = Q_{M-Burden} + \frac{\left(\sum_{i}^{k} U_{i}^{2} - \sum_{i}^{k} V_{i}\right)^{2}}{2\sum_{i}^{k} V_{i}^{2}}$$

For the VT test we test:

$$Q_{RM-VT} = \max_{p} Q_{RM-\text{Burden}}(p),$$

which is the maximum of the random-effects burden tests obtained p[51]. 368 with a MAF threshold p[51].

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SKAT and SKAT-0In order to obtain the test statistic for the SKAT meta-analysis we371assume that the mean of the genetic effects of the variants in each372region, for each study, is drawn from the distribution:373

$$\beta_i = \mu + \xi_i, \quad i = 1, \dots, k$$

where  $\mu$  is the mean of the effects of the *j* variants across the 374 *k* studies and  $\xi_i$  is the deviation of the effects of the *k*<sup>th</sup> study from 375 the mean effect [34, 51]. Random effects models have been pro-376 posed for the SKAT and SKAT-O tests, by Lee et al. [47], the 377 Het-SKAT and Het-SKAT-O and by Tang et al. [51], the 378 RE-SKAT and RE-SKAT-O. RE-SKAT aims to detect mean effects 379 and heterogeneity and HET-SKAT aims to detect heterogeneity in 380 the absence of mean effects [46]. 381

Tang and Lin [46] performed extensive simulation testing to 382 compare the power of the different statistical models. They com- 383 pared three different genetic structure models: rare variant model 384 (for this model it is assumed that 50% of the variants with 385 MAF < 0.5% are causal), low-frequency-variant model (where 386 50% of all variants are assumed to be causal), and opposite effects 387 model (where 50% of the variants are assumed to be causal, 80% of 388 the causal variants are risk, and 20% of the causal variants are 389 protective) [46]. Under these genetic models they evaluated the 390 different meta-analysis methods for the fixed- and random-effects 391 burden, VT, SKAT, and SKAT-O models using normalized score 392 statistics [46]. The tests were performed assuming that the effect of 393 each study is a random variable with mean  $\mu$  and variance  $\tau$ . Two 394 different mean effect values,  $\mu = 0$  and  $\mu = 0.25$ , were tested and 395 the variance  $\tau$  was allowed to vary between 0 and 0.25. For  $\mu = 0.25$  396 and when genetic heterogeneity was small the fixed-effects burden 397 and VT models were more powerful than their equivalent random- 398 effects models and the fixed-effect SKAT and SKAT-O had similar 399 power to their corresponding random-effects models [46]. In the 400 presence of strong heterogeneity the random-effect models were 401 more powerful to the equivalent fixed-effects. The simulations 402 showed that under the rare variant model VT tests were the most 403 powerful whereas for the opposite effects model SKAT and SKAT- 404 O were more powerful than burden and VT models [46]. Under 405 the opposite effects structure model, the random effects models for 406 SKAT and SKAT-O (HET-SKAT and HET-SKAT-O) proposed by 407 Lee et al. [47] were less powerful compared to random effects 408 models proposed by Tang et al. [51] (RE-SKAT and RE-SKAT- 409 O) when the heterogeneity was low, and slightly more powerful 410 when the heterogeneity was large [46]. Under the assumption that 411  $\mu = 0$ , the random effects models were more powerful in the 412 presence of strong heterogeneity. For the rare and low-frequency 413 variant models RE-SKAT-O was the most powerful whereas for the 414 opposite effects model HET-SKAT and HET-SKAT-O were more 415 powerful [46]. Tang and Lin [46] further compared the different 416 methods using normalized and non-normalized score statistics and 417 illustrated that the use of non-normalized score-type statistics can 418 result in power loss. It is obvious that no single test is more 419

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powerful under all different genetic models and since the underlying genetic model is not known upfront, it is important that the different methods are explored in each different case. 422

#### 423

#### 4 Heterogeneity

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There can be numerous reasons underlying between-study hetero- 425 geneity in genomic studies, including differences due to different 426 populations being studied and environmental or lifestyle factors. 427 Other potential sources of heterogeneity include genotyping 428 errors, differences in genotyping platforms, variants being imputed 429 or genotyped in the different studies, or differences in the defini- 430 tions of the phenotype or trait [2]. Different measures of heteroge- 431 neity have been proposed and used in the literature for assessing the 432 between-study differences in underlying models for the common 433 variants in GWAS. The Cochran's Q-statistic [20] and the  $I^2$  metric 434 [52] have been widely used in common variant meta-analyses. Q 435 statistic is used for testing the hypothesis of no heterogeneity 436 between studies and  $I^2$  is a measure of the proportion of the total 437 variability that is due to heterogeneity and takes values between 438 0 and 100%. Usually, a Q statistic *p*-value of < 0.1 is regarded as 439 significant heterogeneity,  $I^2 > 50\%$  is considered moderate hetero- 440 geneity and  $I^2 > 75\%$  is considered as high heterogeneity [52]. 441

Heterogeneity is expected to play a more significant role in the 442 meta-analyses of rare variant studies as rare variants are population 443 specific [53] and can be more sensitive to sequencing technologies 444 errors, quality control measures, and differences in regional 445 annotations [51].

#### 5 Meta-Analysis Software

The majority of the software developed for the common variant 448 meta-analysis will perform both fixed- and random–effects meta-449 analysis for binary and quantitative traits and also allow for geno-450 typed and imputed variants. For the common variant single variant 451 meta-analysis, METAL [54], META [55], GWAMA [56], PLINK 452 [57] and different R packages [58] like MetABEL [59] have been 453 most widely used. MANTRA [31] and TransMeta [27] have been 454 used for the meta-analysis of results from multiethnic studies and 455 METASOFT for the new random effects model proposed by Han 456 and Eskin [25] and the binary effects assumption [26]. 457 The different software require input of individual study summary 458 statistics including *p*-values, sample size, estimate of the regression 459 coefficient, standard errors, and imputation quality metrics. 460 GWAMA, META, and METASOFT perform both fixed- and 461

random effects meta-analysis whereas METAL, PLINK, and MetA-BEL perform only fixed effects meta-analysis. META and METAL 463 also perform *z*-score-based meta-analysis and METASOFT imple-464 ments two additional random-effects meta-analysis tests; a test that 465 is optimal for detecting associations in the presence of heterogene-466 ity and a test that is optimal when some studies have an effect and 467 others do not. All software allow for genomic control adjustments 468 and also produce measures of heterogeneity across the different 469 studies (Cochran's Q *p*-value and  $I^2$ ). 470

Different software have been developed for the meta-analysis of 471 the regional rare variant test statistics including meta-analysis of 472 score statistics (MASS) [50], RAREMETAL [60], MAGA and 473 different R packages like MetaSKAT [47] and segMeta [48]. Each 474 package has their own function or complementary software for the 475 calculation of the score statistics and covariance matrices of each 476 individual study to be used subsequently for the meta-analysis. 477 MASS, MetaSKAT, and seqMeta can be used to obtain summary 478 statistics for both quantitative and binary traits whereas RAREME- 479 TAL can currently only be used for quantitative traits. seqMETA 480 can also be used for survival analysis data and allows for different 481 selection weights for the Burden and SKAT part of the SKAT-O. 482 RAREMETAL and seqMeta also support the analysis of family data 483 and conditional analyses. Furthermore, Tang and Lin developed a 484 software to convert the summary statistics of the different rare 485 variant meta-analysis software (PreMeta) as they are not always 486 compatible [46]. This allows for the easier exchange of summary 487 level statistics across the different studies without the need for each 488 analyst to perform the analysis using the same software. Tang and 489 Lin [46] have also proposed different transformation methods that 490 are implemented in PreMeta which aim to achieve normality and 491 reduce the type I error; the inverse-normal transformation (INT) 492 and rescaled INT (R-INT). MAGA allows for the reconstruction of 493 the regional test statistic using the single variant results from each 494 study (effect estimates and standard errors) and uses correlation 495 matrices from one of the component studies or publicly available 496 resources [49]. The method implemented in MAGA is attractive 497 when each of the component studies cannot obtain the score 498 statistics and information matrices, but has limitation due to the 499 fact that rare variants for which the effect cannot be estimated are 500 not used in the analysis [49]. 501

#### 6 Discussion

Meta-analysis provides a powerful tool for the combination of the 503 results of different studies in order to identify associations that 504 would have not been found through a single study. This has proven 505 an extremely successful method in GWAS and aided the 506

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identification of thousands of robustly associated common variants 507 with complex traits. Although different meta-analysis techniques 508 have been proposed, the fixed-effects meta-analysis has been most 509 widely used in GWAS. The field of genetics is currently being driven 510 by rare variant studies and a large number of sequencing experi-511 ments are currently being performed. Larger power will need to be 512 achieved in order to identify rare variant associations and aggrega-513 tion methods and meta-analysis will clearly play an important role 514 in this identification. As the underlying genetic effects of rare-515 variants are not known upfront and currently there is not a univer-516 sally more powerful rare variant meta-analysis method, a collection 517 of approaches need to be explored. Further validation will need to 518 be performed to regions identified through meta-analyses of aggregation tests.

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Author's Proof

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AU5

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