

Lab session 8
Matched case-control studies
Conditional logistic regression

Open the “endom_ca.dta” file and list the first 20 records. Notice the structure of the dataset and the coding of groups along with the coding of the four controls per group

```
. sort id conno
. li id casecon conno age estrog in 1/20
```

| | id | casecon | conno | age | estrog |
|-----|----|---------|-------|-----|--------|
| 1. | 1 | Case | 0 | 74 | Yes |
| 2. | 1 | Control | 1 | 75 | No |
| 3. | 1 | Control | 2 | 74 | No |
| 4. | 1 | Control | 3 | 74 | No |
| 5. | 1 | Control | 4 | 75 | Yes |
| 6. | 2 | Case | 0 | 67 | Yes |
| 7. | 2 | Control | 1 | 67 | Yes |
| 8. | 2 | Control | 2 | 67 | No |
| 9. | 2 | Control | 3 | 67 | Yes |
| 10. | 2 | Control | 4 | 68 | Yes |
| 11. | 3 | Case | 0 | 76 | Yes |
| 12. | 3 | Control | 1 | 76 | Yes |
| 13. | 3 | Control | 2 | 76 | Yes |
| 14. | 3 | Control | 3 | 76 | Yes |
| 15. | 3 | Control | 4 | 77 | Yes |
| 16. | 4 | Case | 0 | 71 | Yes |
| 17. | 4 | Control | 1 | 70 | Yes |
| 18. | 4 | Control | 2 | 70 | Yes |
| 19. | 4 | Control | 3 | 71 | Yes |
| 20. | 4 | Control | 4 | 70 | Yes |

This is a 1:4 matched case-control study but for a minute we will pretend that the matching scheme is 1:1. Thus we will keep only the first in each group.

```
. drop if conno>1
(189 observations deleted)
. li id casecon conno age estrog in 1/8
```

| | id | casecon | conno | age | estrog |
|----|----|---------|-------|-----|--------|
| 1. | 1 | Case | 0 | 74 | Yes |
| 2. | 1 | Control | 1 | 75 | No |
| 3. | 2 | Case | 0 | 67 | Yes |
| 4. | 2 | Control | 1 | 67 | Yes |
| 5. | 3 | Case | 0 | 76 | Yes |
| 6. | 3 | Control | 1 | 76 | Yes |
| 7. | 4 | Case | 0 | 71 | Yes |
| 8. | 4 | Control | 1 | 70 | Yes |

To analyze this dataset using Mc Nemar’s test (paired χ^2) data need to be in wide instead of long format. This can be easily obtained using Stata’s reshape command.

```
. reshape wide age gall hyper obesity estrog dose dur nonestr conno, i(id) j(casecon)
> secon)
(note: j = 0 1)
```

```
Data                long   ->   wide
-----
Number of obs.      126   ->    63
Number of variables  11   ->    19
j variable (2 values) casecon -> (dropped)
xij variables:
      age   ->  age0 age1
      gall ->  gall0 gall1
      hyper ->  hyper0 hyper1
      obesity ->  obesity0 obesity1
      estrog ->  estrog0 estrog1
      dose ->  dose0 dose1
      dur ->  dur0 dur1
      nonestr ->  nonestr0 nonestr1
      conno ->  conno0 conno1
-----
```

Now list the first 4 records to check the wide format and compare with the previous listing

```
. li id estrog* age* in 1/4
```

| | id | estrog0 | estrog1 | age0 | age1 |
|----|----|---------|---------|------|------|
| 1. | 1 | No | Yes | 75 | 74 |
| 2. | 2 | Yes | Yes | 67 | 67 |
| 3. | 3 | Yes | Yes | 76 | 76 |
| 4. | 4 | Yes | Yes | 70 | 71 |

Data are now in a suitable format to perform the Mc Nemar's test using Stata's `mcc` command. We first check the effect of oestrogen use on the risk of endometrial cancer

```
. mcc estrog1 estrog0
```

| Cases | Controls | | Total |
|-----------|----------|-----------|-------|
| | Exposed | Unexposed | |
| Exposed | 27 | 29 | 56 |
| Unexposed | 3 | 4 | 7 |
| Total | 30 | 33 | 63 |

```
McNemar's chi2(1) = 21.12 Pr>chi2 = 0.0000
Exact McNemar significance probability = 0.0000
```

Proportion with factor

| | Cases | Controls | [95% conf. interval] | |
|------------|----------|----------|----------------------|---------|
| difference | .4126984 | .253346 | .5720509 | |
| ratio | 1.866667 | 1.424262 | 2.446492 | |
| rel. diff. | .7878788 | .6331393 | .9426183 | |
| odds ratio | 9.666667 | 2.996311 | 49.58254 | (exact) |

Similarly we can test for the effect of gall-bladder disease

```
. mcc gall1 gall0
```

| Cases | Controls | | Total |
|-----------|----------|-----------|-------|
| | Exposed | Unexposed | |
| Exposed | 4 | 13 | 17 |
| Unexposed | 5 | 41 | 46 |
| Total | 9 | 54 | 63 |

McNemar's chi2(1) = 3.56 Pr>chi2 = 0.0593
Exact McNemar significance probability = 0.0963

Proportion with factor

| | Cases | Controls | [95% conf. interval] | |
|------------|----------|----------|----------------------|------------------|
| difference | .1269841 | | -.017101 | .2710693 |
| ratio | 1.888889 | | .9643767 | 3.699697 |
| rel. diff. | .1481481 | | .0060224 | .2902738 |
| odds ratio | 2.6 | | .8698097 | 9.315215 (exact) |

a) How is the Odds Ratio and the χ^2 statistic calculated?

Dividing the pairs according to the age of the case (under 70 – 70 and older) we find

```
. gen age70=age1>=70
```

Under 70 years

```
mcc gall1 gall0 if age70==0
```

| Cases | Controls | | Total |
|-----------|----------|-----------|-------|
| | Exposed | Unexposed | |
| Exposed | 2 | 7 | 9 |
| Unexposed | 1 | 18 | 19 |
| Total | 3 | 25 | 28 |

McNemar's chi2(1) = 4.50 Pr>chi2 = 0.0339
Exact McNemar significance probability = 0.0703

Proportion with factor

| | Cases | Controls | [95% conf. interval] | |
|------------|----------|----------|----------------------|-----------------|
| difference | .2142857 | | -.0028088 | .4313802 |
| ratio | 3 | | 1.032252 | 8.718799 |
| rel. diff. | .24 | | .0466875 | .4333125 |
| odds ratio | 7 | | .8993189 | 315.599 (exact) |

70 years and older

```
. mcc gall1 gall0 if age70==1
```

| Cases | Controls | | Total |
|-----------|----------|-----------|-------|
| | Exposed | Unexposed | |
| Exposed | 2 | 6 | 8 |
| Unexposed | 4 | 23 | 27 |
| Total | 6 | 29 | 35 |

McNemar's chi2(1) = 0.40 Pr>chi2 = 0.5271
Exact McNemar significance probability = 0.7539

Proportion with factor

| | Cases | Controls | [95% conf. interval] | |
|------------|----------|----------|----------------------|------------------|
| difference | .2285714 | .1714286 | -.147498 | .2617838 |
| ratio | 1.333333 | | .5450297 | 3.2618 |
| rel. diff. | .0689655 | | -.1372555 | .2751866 |
| odds ratio | 1.5 | | .3557302 | 7.226552 (exact) |

To test for the homogeneity of the two OR's in the different age groups we can form the following 2x2 table and calculate the usual χ^2 statistic or most preferably the Fisher's exact significance.

```
. cci 7 6 1 4,exact
```

| | Exposed | Unexposed | Total | Proportion Exposed |
|----------|---------|-----------|-------|--------------------|
| Cases | 7 | 6 | 13 | 0.5385 |
| Controls | 1 | 4 | 5 | 0.2000 |
| Total | 8 | 10 | 18 | 0.4444 |

Point estimate [95% Conf. Interval]

| | | | |
|-----------------|----------|-----------|-------------|
| Odds ratio | 4.666667 | .5013338 | (Cornfield) |
| Attr. frac. ex. | .7857143 | -.9946789 | (Cornfield) |
| Attr. frac. pop | .4230769 | | |

1-sided Fisher's exact P = 0.2255
2-sided Fisher's exact P = 0.3137

Thus there is no evidence for a modifying effect of age on the OR for gall-bladder disease.

If we try to evaluate hypertensive disease as a confounding or modifying factor in similar fashion, we find that there is severe loss of data because of the restriction to case-control pairs, which are homogeneous for hypertension

```
. count
63
. count if hyper1==hyper0
32
. di (1-32/63)*100
49.206349
```



```

. use endom_ca,clear

. clogit casecon estrogen,group(id) nolog or

Conditional (fixed-effects) logistic regression   Number of obs   =       315
                                                    LR chi2(1)      =       35.35
                                                    Prob > chi2     =       0.0000
Log likelihood = -83.72159                       Pseudo R2       =       0.1743

```

| casecon | Odds Ratio | Std. Err. | Z | P> z | [95% Conf. Interval] |
|----------|------------|-----------|-------|-------|----------------------|
| estrogen | 7.954681 | 3.347525 | 4.928 | 0.000 | 3.486714 18.14802 |

```

. est store A

```

Notice the reduced estimate for the OR along with the reduced value of its Standard error.

c) How is the “LR chi2(1)=35.35” statistic calculated? What is the meaning of this result?

Interactions of estrogens with age

While the main effects of age cannot be tested (matched variable) interactions of estrogen with age CAN BE TESTED

Before we proceed we must create dummy variables for two of the three age groups (1:55-64, 2:65-74, 3:75+) and their interactions with the “estrog” variable. We will use as age for controls the age of the corresponding control pretending that we have exact age matching (in fact differences do not exceed 2 yrs).

```

. gsort id -casecon
. qui by id:gen age32=age[1]>=65 & age[1]<=74
. qui by id:gen age33=age[1]>=75
. gen age32est=age32*estrog
. gen age33est=age33*estrog

```

Now we will fit the model:

```

. clogit casecon estrogen age32est age33est,group(id)

Conditional (fixed-effects) logistic regression   Number of obs   =       315
                                                    LR chi2(3)      =       36.03
                                                    Prob > chi2     =       0.0000
Log likelihood = -83.380155                       Pseudo R2       =       0.1777

```

| casecon | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|----------|----------|-----------|------|-------|----------------------|
| estrogen | 1.430827 | .8256893 | 1.73 | 0.083 | -.1874939 3.049149 |
| age32est | .847401 | 1.03377 | 0.82 | 0.412 | -1.178751 2.873553 |
| age33est | .7801409 | 1.154229 | 0.68 | 0.499 | -1.482107 3.042389 |

```

. est store B

```

We can check the significance of the interaction as follows:

```
. lrtest A B

likelihood-ratio test          LR chi2(2)  =      0.68
(Assumption: A nested in B)   Prob > chi2 =      0.7107
```

The p=0.71 indicating that the differences by age group ARE not statistically significant. Therefore separate OR's by age groups should not be reported.

Alternative way of doing the same thing

```
gsort id -casecon
by id:gen case_age=age[1]
egen agegr=cut(case_age),at(0,65,75,200) label
xi i.estrog*i.agegr
clogit casecon _Iestrog_1 _IestXage_1_1 _IestXage_1_2,group(id)
```

Other covariates: Gall-blaster disease

```
. clogit casecon estrog gall,group(id) nolog

Conditional (fixed-effects) logistic regression   Number of obs   =      315
                                                  LR chi2(2)      =      45.05
                                                  Prob > chi2     =      0.0000
Log likelihood = -78.871308                    Pseudo R2       =      0.2221

-----+-----
casecon |          Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
  estrog |   2.114785     .439794     4.809  0.000     1.252804   2.976765
   gall  |   1.274654     .4108678    3.102  0.002     .4693684   2.079941
-----+-----

. est store C
```

Gall disease is a significant predictor of endometrial cancer:

OR:exp(1.274654)=3.58. That is, women with Gall disease have 3.58 (95% CI: 1.59 – 8.0) times higher probability (odds) to develop endometrial cancer than women without Gall disease.

According to the Wald test: P=0.002. The OR for estrogens has not been substantially changed (OR=8.29; 95% CI: 3.50–19.62).

Interactions between estrogens and Gall disease

We can check for interaction between gall-bladder disease and oestrogen use

```
. gen estgall=estrog*gall

. clogit casecon estrog gall estgall,group(id) nolog

Conditional (fixed-effects) logistic regression   Number of obs   =      315
                                                  LR chi2(3)      =      49.33
                                                  Prob > chi2     =      0.0000
Log likelihood = -76.730576                    Pseudo R2       =      0.2432

-----+-----
casecon |          Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
  estrog |   2.700139     .6117687    4.414  0.000     1.501094   3.899183
   gall  |   2.894345     .883053     3.278  0.001     1.163593   4.625097
 estgall |  -2.052747     .9949737   -2.063  0.039    -4.002859  -1.1026342
```

```
-----
. est store D
-----
```

According to the Wald test interaction is significant (P=0.039). **NOTE: We have negative interaction** (i.e., the interaction term is negative)
We can also use LR-test

```
. lrtest C D
```

```
likelihood-ratio test                LR chi2(1)  =      4.28
(Assumption: C nested in D)         Prob > chi2 =    0.0385
```

The two tests give almost identical results

d) Fill the table below. Are the effects of estrogen use more likely to be additively combined or multiplicatively with those of Gall disease?

Report of interactions

Estrogens

| | | Yes | No |
|---------|-----|-----|----|
| Gall | Yes | | |
| Disease | No | | |

We can get the same results using the lincom command and then exponentiating the results

```
. lincom estrog
( 1)  estrog = 0.0
-----
casecon |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      (1) |  2.700139   .6117687    4.414  0.000     1.501094    3.899183
-----
. lincom gall
( 1)  gall = 0.0
-----
casecon |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      (1) |  2.894345   .883053    3.278  0.001     1.163593    4.625097
-----
. lincom gall+estrog+estgall
( 1)  estrog + gall + estgall = 0.0
-----
casecon |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      (1) |  3.541737   .7232228    4.897  0.000     2.124246    4.959227
-----
```

```
. di "[OR (95% CI) | estrog=Yes, gall=No] = ",exp( 2.700139 ),"("
exp(1.501094), ", "exp(3.899183),")"
[OR (95% CI) | estrog=Yes, gall=No] =  14.8818 (4.4865947 , 49.362104 )

. di "[OR (95% CI) | estrog=No, gall =Yes]= ",exp( 2.894345 ),"(" exp(1.163593
) ,", "exp(4.625097 ),")"
[OR (95% CI) | estrog=No, gall=Yes] =  18.071661 (3.2014153 , 102.01267 )

. di "[OR (95% CI) | estrog=Yes, gall=Yes]= ",exp(3.541737 ),"(" exp( 2.124246
) ,", "exp(4.959227 ),")"
```


[OR (95% CI) | estrog=Yes, gall=Yes]= 34.52684 (8.3665867 , 142.48361)