

Μάθημα Ασκήσεων (S. Poisson κ. Ανα. Θεωρία)

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Άσκηση 3.6 (Θεμα 1. Ιουνίου 2018)

$N_1(t), N_2(t)$  αριθμοί S. Poisson με ρυθμούς  $\lambda, \mu$ .

$$N(t) = N_1(t) + N_2(t) \text{ υπόρθεση}$$

1.  $P(\text{Το } 1^{\text{ο}} \text{ γεγονός της } N_1 \text{ να είναι το } n^{\text{ο}} \text{ γεγονός της } N)$

↳ Εάντο  $S_1^{(n)}$  ο χρόνος του  $k^{\text{ο}}$  γεγονός της  $N_i$ ,  $i=1,2$ .

[ $\Sigma_{t_0} \omega_{1,t}$ ] να εκπονεί η γεγ. ωρά την  $N(t)$  και

1 γεγ. φέρει την  $N_1$ ,

$\Leftrightarrow$

$\Sigma_{t_0} \omega_{1,t}$  να εκπονεί 1 γεγονός  $N_1$  και  $n-1$   $N_2$

Συντομεύσεις

$$\begin{aligned} P(N_2(S_1^{(n)})=n-1) &= \int_0^\infty P(N_2(S_1^{(n)})=n-1 | S_1^{(n)}=x) f_{S_1^{(n)}}(x) dx \\ S_1^{(n)} &\sim \text{Exp}(\lambda) \\ &= \int_0^\infty P(N_2(x)=n-1) f_{S_1}(x) dx \\ &= \int_0^\infty e^{-\mu x} \frac{(\mu x)^{n-1}}{(n-1)!} \lambda e^{-\lambda x} \lambda e^{-\lambda x} dx \\ &= \underbrace{\frac{\mu^{n-1} \lambda}{(\lambda+\mu)^n} \int_0^\infty (\lambda+\mu)^n \frac{x^{n-1}}{(n-1)!} e^{-(\lambda+\mu)x} dx}_{\text{Erlang}(n, \lambda+\mu)} \\ &= \left( \frac{\mu}{\lambda+\mu} \right)^{n-1} \frac{\lambda}{\lambda+\mu} \end{aligned}$$

α γράμμως

$$\begin{aligned} \text{Επίσημο } p &= P(\text{Το } \sigma\text{ γεγονός της } N(t) \text{ να είναι } \omega \text{ τη } N_2(t)) \\ &= \underbrace{P(S_1^{(2)} < S_1^{(n)})}_{\text{ }} = \frac{\mu}{\lambda+\mu} \end{aligned}$$

Επίσημο  $\{I_i=j\} = \{\text{Το } i^{\text{ο}} \text{ γεγονός είναι } \omega \text{ τη } N_j\}$ ,  $j=1,2$

$$\begin{aligned} \text{Znacíme m.v. } & \Pr(I_1=2, I_2=2, \dots, I_{n-1}=2, I_n=1) \\ & = \Pr(I_1=2)^{n-1} \Pr(I_n=1) \\ & = \left(\frac{\mu}{\lambda+\mu}\right)^{n-1} \frac{1}{\lambda+\mu} \end{aligned}$$

→

2 Na určujeme n  $\Pr(N_1(t)=1 \mid N(t)=n)$

$$\begin{aligned} \hookrightarrow \Pr(N_1(t)=1 \mid N(t)=n) &= \frac{\Pr(N_1(t)=1, N_2(t)+N_3(t)=n)}{\Pr(N(t)=n)} \\ &= \frac{\Pr(N_1(t)=1, N_2(t)=n-1)}{\Pr(N(t)=n)} \stackrel{N_1, N_2 \text{ a.v.s.}}{=} \frac{\Pr(N_1(t)=1) \Pr(N_2(t)=n-1)}{\Pr(N(t)=n)} \quad (1) \end{aligned}$$

$N_1 \sim P(\lambda t), N_2 \sim P(\mu t), N \sim P((\lambda+\mu)t)$

$$(1) = \frac{e^{-\lambda t} \frac{\lambda t}{\lambda} \cdot e^{-\mu t} \frac{(\mu t)^{n-1}}{(n-1)!}}{e^{-\frac{(\lambda+\mu)t}{\lambda+\mu} + (\lambda+\mu)t} \frac{n!}{n!}} = \frac{1}{\lambda+\mu} \cdot \left(\frac{\mu}{\lambda+\mu}\right)^{n-1}$$

→

3.  $\Pr(N(t)=n \mid N_2(t/2)=n-1) = j$

$$\hookrightarrow \frac{\Pr(N_1(t)+N_2(t)=n, N_2(t/2)=n-1)}{\Pr(N_2(t/2)=n-1)}$$

$$= \frac{\Pr(\underbrace{N_1(t)}_{N_1(t)}, \underbrace{N_2(t)-N_2(t/2)}_{N_2(t/2)}=1, N_2(t/2)=n-1)}{\Pr(N_2(t/2)=n-1)} \quad (2)$$

$N_2(t) - N_2(t/2), N_2(t/2)$  a.v.s.  $\Rightarrow \begin{cases} N_1(t) + N_2(t) - N_2(t/2) \\ N_2(t/2) \end{cases}$  a.v.s.

$$(2) = \frac{\Pr(N_1(t) + N_2(t) - N_2(t/2) = 1)}{\Pr(N_2(t/2) = n-1)} \Pr(N_2(t/2) = n-1)$$

$$= \Pr(\underbrace{N_1(t) + N_2(t) - N_2(t/2)}_{\sim P((\lambda+\frac{\mu}{2})t)} = 1) \xrightarrow{\text{ano co vypočítal}} \Pr((\lambda+\frac{\mu}{2})t)$$

$$= P(N_1(t) = 1) P(N_2(t/2) = 0) + P(N_1(t) = 0) P(N_2(t/2) = 1)$$

$$= e^{-\lambda t} \lambda t e^{-\mu t/2} + e^{-\lambda t} e^{-\mu t/2} \mu t/2$$


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4. Συνοπε των πιθανότητων της  $N(t) - N_2(t/2)$

$$\hookrightarrow \underbrace{N_1(t)}_{P(\lambda t)} + \underbrace{N_2(t) - N_2(t/2)}_{P(\mu t/2)} \quad \text{ανο δεμορία}$$

$$0 < t_1 < t_2$$

$$N(t_1+t_2) - N(t_1) \xrightarrow{\text{τοσκ.}} N(t_2)$$

ανο δεμορία (αρχιποληπτικά Poisson  
→ Poisson )

Άρα  $N_1(t) + N_2(t) - N_2(t/2) \sim P(\lambda t + \mu t/2)$

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$$\begin{aligned} 5. \text{Cov}(N_1(t), N(t)) &= E[N_1(t) N(t)] - E[N_1(t)] E[N(t)] \\ &= E[N_1^2(t) + N_1(t) N_2(t)] - \lambda t (\lambda + \mu) t \\ &= E[N_1^2(t)] + E[N_1(t) N_2(t)] - \lambda t (\lambda + \mu) t \\ &= E[N_1^2(t)] - E[N_1(t)]^2 + E[N_1(t)]^2 \\ &\quad + \cancel{\lambda \mu t^2} - \cancel{\lambda^2 t^2} - \cancel{\lambda \mu t^2} \\ &= V[N(t)] + \cancel{(\lambda t)^2} - \cancel{(\mu t)^2} \\ &= V[N(t)] = \lambda t \end{aligned}$$

αλγαρίδες (κειδούρες συνδεσμ.)

$$\begin{aligned} \text{Cov}(N_1(t), N_1(t) + N_2(t)) &= \text{Cov}(N_1(t), N_1(t)) \\ &\quad + \text{Cov}(N_1(t), N_2(t)) \quad N_2, N_1 \text{ ανο} \\ &= V[N_1(t)] = \lambda t \end{aligned}$$


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### Ασκηση 3.5 (Θερινα 1 Σεπτ 2019)

ενη θεση  $N_1 \sim P(\lambda)$ ,  $\lambda = 5/\text{ωρα}$  παραγκ  $N_2 \sim P(\mu)$ ,  $\mu = 25/\text{ωρα}$

1.  $E \subseteq \underbrace{N_2(1) \mid N(1)=60}_\text{Bin(n, p)}$

$$n=60 \quad p = \frac{\mu}{\lambda+\mu} = \frac{25}{30}$$

Άρα

$$E = [N_2(1) \mid N(1)=60] =$$

$$np = 50$$

$$N(t) = N_1(t) + N_2(t) \sim P(\lambda+\mu)t$$

$$N(t) = N_1(t) + N_2(t) + \dots + N_k(t)$$

$$N_j(t) \mid N(t)=n$$

$$\sim Bin(n, p_j)$$

$$p_j = \frac{\lambda_j}{\sum \lambda_i}$$

2.  $P(N_1(1)=10 \mid N_1(2)=15, \underbrace{N_2(2)=60})$

$$= \frac{P(N_1(1)=10, N_1(2)=15, N_2(2)=60)}{P(N_1(2)=15, N_2(2)=60)}$$

$$= \frac{P(N_1(1)=10, N_1(2)=15)}{P(N_1(2)=15)}$$

$$= \frac{P(N_1(1)=10, N_1(2)-N_1(1)=5)}{P(N_1(2)=15)}$$

$$= \frac{P(N_1(1)=10) P(N_1(2)-N_1(1)=5)}{P(N_1(2)=15)} = \dots$$

3. Οπως συνεχ 3.6 το 1.

$$P(I_1=2, I_2=2, I_3=2, I_4=2, I_5=1) = \left(\frac{\lambda}{\lambda+\mu}\right)^4 \frac{\lambda}{\lambda+\mu}$$

4.  $P(N_2(2) = 100 \mid N_2(1) = 70, N_1(1) = 35)$   
 αντως το ερώτημα 2.

5. ( $S_R$ ) χρονοί γεγονότων της  $N(t)$ ,  $S_R \sim \text{Exponential}(\kappa, t + \mu)$

Ζητάμε την

$$E[S_{10} \mid N(1) = 29] = E[U_{10, 29}] = \frac{10}{30} = 20 \text{ min.}$$

$$\kappa \leq n: E[S_{10} \mid N(1) = 29] = \frac{\kappa t}{n+1}$$

$$E[S_R \mid N(t) = n]$$

$$E[S_{34} \mid N(1) = 29] = 1 + 5E[S_1] = 1 + \frac{5}{30} = \frac{35}{30}$$

Γενικά, για  $\underline{k \geq n}$

$$E[S_k \mid N(t) = n] = t + (k-n)E[S_1]$$

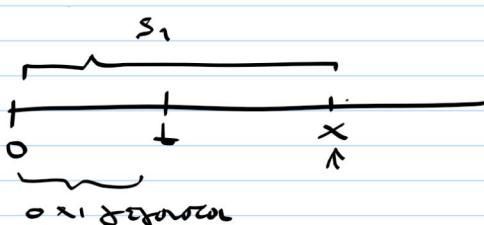
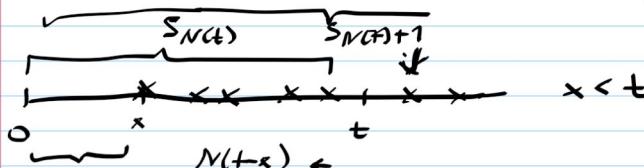
Άσκηση 2.6 (Θεμα 2 Ιανουαρίου 2018)

$$h(t) = E[\frac{S_{N(t)} + S_{N(t)+1}}{2}]$$

1. Αναν. εξιών.

2. Ράσος τυχών για τη θύμη στα  $N(t) \sim P(\lambda t)$

↪ 1: Δεσμεύουμε στο χρονό του 1<sup>ου</sup> γεγ. δοθεί  $S_1 = x$



$$S_{N(t)} / S_1 = x = \begin{cases} 0 & x \neq t \\ x + S_{N(t-x)} & x \leq t \end{cases}$$

$$S_{N(t)+1} / S_1 = x = \begin{cases} x & x > t \\ x + S_{N(t-x)+1} & x \leq t \end{cases}$$

$$\begin{aligned}
 \text{Apa} \quad h(t) &= E \left[ \frac{S_{N(t)} + S_{N(t)+1}}{2} \right] = \\
 &= \int_0^t E \left[ \frac{x + S_{N(t-x)} + x + S_{N(t-x)+1}}{2} \right] dF_{S_1}(x) \\
 &+ \int_t^\infty \frac{x}{2} dF_{S_1}(x) = \int_0^t x dF_{S_1}(x) + \int_t^\infty \frac{x}{2} dF_{S_1}(x) \\
 &+ \int_0^t E \left[ \frac{S_{N(t-x)} + S_{N(t-x)+1}}{2} \right] dF_{S_1}(t) \\
 &\qquad\qquad\qquad \underbrace{\qquad\qquad\qquad}_{h(t-x)} \\
 &= D(t) + \int_0^t h(t-x) dF_{S_1}(t) \\
 \text{per } D(t) &= \int_0^t \frac{x}{2} dF_{S_1}(x) + \frac{1}{2} \underbrace{\int_0^\infty x dF_{S_1}(x)}_{E[S_1]} \quad (3)
 \end{aligned}$$

2. H. Apon tns Brav.  $\Sigma$  follows (3) eva.

$$\begin{aligned}
 h(t) &= D(t) + \int_0^t D(t-x) dm(x) \\
 \text{orwo } D(t) &= \frac{1}{2} \int_0^t x dF_{S_1}(x) + \frac{1}{2} E[S_1]
 \end{aligned}$$

$$\text{Kai n } m(t) = E[N(t)]$$

$\Sigma$   $\delta$  exposure or,  $N(t) \sim P(\lambda t)$  kai apa  
 $S_1 \sim \text{Exp}(\lambda)$

Once,

$$\begin{aligned}
 D(t) &= \frac{1}{2} \int_0^t x \lambda e^{-\lambda x} dx + \frac{1}{2\lambda} \\
 &= \frac{1}{2} \left[ -x e^{-\lambda x} \Big|_0^t - \frac{e^{-\lambda x}}{\lambda} \Big|_0^t \right] + \frac{1}{2\lambda} \\
 &= \frac{1}{2} \left[ -t e^{-\lambda t} - \frac{e^{-\lambda t}}{\lambda} + \frac{1}{\lambda} \right] + \frac{1}{2\lambda}
 \end{aligned}$$

$$= \frac{1}{2} \left[ \frac{2}{\lambda} - e^{-\lambda t} \left( t + \frac{1}{\lambda} \right) \right]$$

$$dm(x) = \lambda dx$$

$\rightarrow \lambda x$

$$h(t) = D(t) + \lambda \int_0^t D(t-x) dx = \begin{aligned} & \stackrel{u=t-x}{=} \lambda \int_0^t D(u) du + D(t) \\ & \stackrel{du=-dx}{=} \\ & \stackrel{u_1=t}{=} \\ & \stackrel{u_2=0}{=} \end{aligned}$$

$$\begin{aligned} \cdot \lambda \int_0^t D(u) du &= \lambda \int_0^t \frac{1}{2} \left[ \frac{2}{\lambda} - e^{-\lambda u} \left( u + \frac{1}{\lambda} \right) \right] du \\ &= \lambda \left[ \frac{1}{\lambda} t - \frac{1}{2} \int_0^t u e^{-\lambda u} - \frac{1}{2} \frac{1}{\lambda^2} \int_0^t e^{-\lambda u} du \right] \\ &= \lambda \left[ \frac{1}{\lambda} t - \frac{1}{2} \left[ -\frac{u e^{-\lambda u}}{\lambda} \Big|_0^t + \frac{e^{-\lambda u}}{\lambda^2} \Big|_0^t \right] - \frac{1}{2\lambda^2} (1 - e^{-\lambda t}) \right] \\ &= t + \frac{1}{2} [t e^{-\lambda t}] + \frac{1}{2\lambda} (e^{-\lambda t} - 1) - \frac{1}{2\lambda} + \frac{e^{-\lambda t}}{2\lambda} \end{aligned}$$

Apa

$$\begin{aligned} h(t) &= \frac{1}{\lambda} \left[ -t \frac{-e^{-\lambda t}}{2} - \frac{e^{-\lambda t}}{2\lambda} \right] + t + t \frac{-e^{-\lambda t}}{2} + \frac{e^{-\lambda t}}{2\lambda} \\ &- \frac{1}{\lambda} + \frac{e^{-\lambda t}}{2\lambda} = t + \frac{e^{-\lambda t}}{2\lambda} \end{aligned}$$


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Ergebnis Verwendung:  $N(t) \sim \text{Poisson } (\lambda)$

$$E[S_{N(t)}] = t - \frac{1 - e^{-\lambda t}}{\lambda} \rightarrow \text{An. Av. EJ: } h(t) = D(t) + \lambda \int_0^t D(t-x) dx$$

$$D(t) = \int_0^t x \cdot dF_x(x)$$

$$E[S_{N(t)+K}] = E[X](m(t) + K)$$

$$\begin{aligned} K \geq 1 &= \frac{1}{\lambda} (\lambda t + K) = t + \frac{K}{\lambda} \end{aligned}$$

$$E[A(t)] = \frac{1 - e^{-\lambda t}}{\lambda}$$

$\uparrow$   
 $t - S_{N(t)}$

$$E[B(t)] = \frac{1}{\lambda}$$

$\uparrow$   
 $S_{N(t)+1} - t$

$$E[C(t)] = \frac{1}{\lambda} (2 - e^{-\lambda t}) \quad \rightarrow \text{Núm av. ej: } h(t) = P(t) + \lambda \int_0^t D(t-x) dx$$

$\uparrow$   
 $S_{N(t)+1} - S_{N(t)}$

$h(t) = P(t) + \lambda \int_0^t D(t-x) dx$   
 $D(t) = \int_t^\infty x dF_x(x)$   
 $= t e^{-\lambda t} + \frac{1}{\lambda} e^{-\lambda t}$