Statistical Methods in Epidemiology Lab2 - Solutions. Rates in Follow-up Studies

I. The Diet Data Set.

1. . desc

Contains data from C:\diet.dta
obs: 337 Diet data with dates
vars: 13 25 Sep 2006 16:31
size: 17,187 (99.8% of memory free)

variable name	_	display format	value label	variable label
id doe dox chd dob job month energy height weight fat fibre hieng	long long float long int byte float float float float float	%9.0g %dDmCY %dDmCY %9.0g %dDmCY %8.0g %8.0g %9.0g %9.0g %9.0g %9.0g %9.0g		Subject identity number Date of entry Date of exit Outcome: 1= chd, 0 otherwise Date of birth Occupation month of survey Total energy (1000kcals/day) Height (cm) Weight (kg) Total fat (g/day) Total fibre (g/day) Indicator for energy > 2.75

Sorted by:

. tab chd

Outcome: 1= chd, 0 otherwise	Freq.	Percent	Cum.
0	291 46	86.35 13.65	86.35 100.00
Total	337	100.00	

This table shows the distribution of the outcome of interest, i.e. CHD, within the cohort. There are 13.65% cases with CHD among the 337 subjects.

2. list id doe dox chd in 1/20

_				
	id	doe	dox	chd
1.	1	16Aug1964	01Dec1976	0
2.	2	16Dec1964	01Dec1976	0
3.	3	16Nov1965	01Dec1976	0
4.	4	16Sep1965	01Dec1976	0
5.	5	16Sep1965	31Mar1976	0
6.	 6	16Mar1965	31Aug1968	0
7.	7	16Nov1958	01Dec1976	0
8.	8	16May1965	01Dec1976	0
9.	9	16Feb1959	10Jan1962	0
10.	10	16Jul1964	16May1974	0
		160 11064		
11.	11	160ct1964	08Apr1974	0
12.	12	16Jul1964	03Aug1974	0
13.	13	16Sep1964	16Feb1974	0
14.	14	16Dec1959	01Dec1976	0
15.	15	16May1962	20Aug1976	0
16.	 16	16May1959	31Dec1959	 1
17.	10 17	16Feb1959	14Jan1965	0
18.	17 18	16Feb1959	08Mar1968	0
19.	1 19	16Feb1959	12Mar1966	0
20.	1 20	16Feb1959	26Dec1969	0
	, 20 +			۱ · +

```
337 total obs.

0 exclusions

337 obs. remaining, representing
46 failures in single record/single failure data
4603.669 total analysis time at risk, at risk from t = 0

earliest observed entry t = 0

last observed exit t = 20.04107
```

With this command we specify that the data are survival data, i.e. stset declares data to be survival-time (st) data.

Whenever you type stset or streset, Stata runs or reruns checks on your data making sure that what you are now declaring (or declared in the past) makes sense.

There are 337 total observations with 46 failures. The failures are declared in the fail option. failure(varname[==numlist]) specifies the failure event. failure() must be specified with multiple-record data and is optional with single-record data. If failure() is not specified, every record is assumed to end in a failure. failure(varname) specifies that a failure occurs whenever varname is not zero and not missing. failure(varname[==numlist]) specifies that a failure occurs whenever varname takes on any of the values of numlist.

enter([varname==numlist] time exp) specifies when a subject first comes under observation.

exit(failure | [varname==numlist] time exp) specifies the latest time under which the subject is both under observation and at risk of the failure event.

origin([varname==numlist] time exp|min) and scale(#) define analysis time. origin() defines when a person becomes at risk and scale() can be handy for making t units more readable (such as converting days in years).

4. . li id _t0 _t _d _st in 1/20

-	id id	_t0	_t	_d	_st
1. 2. 3. 4. 5.		0 0 0 0 0	12.29295 11.958932 11.041752 11.208761 10.537988	0 0 0 0 0	1 1 1 1 1
6. 7. 8. 9.	 6 7 8 9 10	0 0 0 0 0	3.4606434 18.042437 11.545517 2.899384 9.8316222	0 0 0 0	1 1 1 1 1
11. 12. 13. 14. 15.	11 12 13 14 15	0 0 0 0	9.4757016 10.047912 9.4182067 16.960986 14.264203	0 0 0 0	1 1 1 1 1
16. 17. 18. 19. 20.	16 17 18 19 20	0 0 0 0	.62696783 5.9110198 9.0568104 7.0663929 10.858316	1 0 0 0	1 1 1 1 1

- _t0: time of entry on the analysis time scale
- _t: time of exit on the analysis time scale
- _d: reason for exit
- _st: indicator for whether the record is included in the analysis.
- 5. . stset dox, fail(chd) origin(dob) enter(doe) scale(365.25)

```
337 obs. remaining, representing
```

```
337 obs. remaining, representing
46 failures in single record/single failure data
4603.669 total analysis time at risk, at risk from t = 0
earliest observed entry t = 30.07529
```

last observed exit t = 69.99863

. li id $_$ t0 $_$ t $_$ d $_$ st in 1/20

_	+				+
	 id	_t0	_t	_d	_st
1.	1	49.615332	61.908282	0	1
2.	2	50.537988	62.49692	0	1
3.	3	58.784394	69.826146	0	1
4.	4	58.726899	69.935661	0	1
5.	5	59.460643	69.998631	0	1
6.	 6	50.98152	54.442163	0	1
7.	7	45.138946	63.181383	0	1
8.	8	50.428474	61.97399	0	1
9.	j 9	67.099247	69.998631	0	1
10.	10	60.167009	69.998631	0	1
11.	 11	60.52293	69.998631	0	1
12.	12	59.950719	69.998631	0	1
13.	13	60.580424	69.998631	0	1
14.	14	43.950719	60.911704	0	1
15.	15	55.734428	69.998631	0	1
16.	 16	 62.661191	63.288159	1	1
17.	17	59.780972	65.691992	0	1
18.	18	60.941821	69.998631	0	1
19.	19	60.960986	68.027379	0	1
20.	20	59.140315	69.998631	0	1
-	+				+

Note that here the subjects are supposed to start participating in the study at _t0 which is the age of entering into the study. We are the ones who specified this with the option origin(dob). The time of exit is the age at exit _t. Note however that the person time at risk is the same as before since it is _t-_t0.

6. tab hieng

Indicator for energy > 2.75	 Freq.	Percent	Cum.
0 1	155 182	45.99 54.01	45.99 100.00
Total	337	100.00	

This is our exposure. 0 denotes subjects with energy intake <=2.75 Mcal and 1 those with energy intake >2.75 Mcal. The table shows the distribution of the exposure among the subjects of the cohort.

7. strate,per(1000)

Estimated rates (per 1000) and lower/upper bounds of 95% confidence intervals (337 records included in the analysis)

+				+
D	Y	Rate	Lower	Upper
46	4.6037	9.9920	7.4843	13.3400

strate tabulates the rate, estimated as the number of failures divided by the person-years, by different levels of one or more categorical explanatory variables (declared in the varlist of the command – see the following command). Confidence intervals for the rate are also calculated.

Here we have 46 chd cases and 4.6 per 1000 person-years in total. The rate is 9.992.

- 8. We generate a new variable for categorizing height:
 - . gen htgrp=height
 (5 missing values generated)
 - . recode htgrp $\min/169.999=0$ 170/174.999=1 175/179.999=2 180/195=3 (htgrp: 332 changes made)
 - . tab htgrp

htgrp	Freq.	Percent	Cum.
0 1 2	92 102 83	27.71 30.72 25.00	27.71 58.43 83.43
3	55	16.57	100.00
Total	332	100.00	

You could check if the recording was successful by trying the following command:

- . sort htgrp
- . by htgrp: sum height, de

._____

-> n	tgrp = 0	Height (c	em)	
	Percentiles	Smallest		
1%	152.4	152.4		
5%	157.988	153.67		
10%	160.02	157.48	Obs	92
25%	163.83	157.734	Sum of Wgt.	92
50%	166.37		Mean	165.5602
		Largest	Std. Dev.	3.530403
75%	167.7924	169.3926		
90%	168.91	169.545	Variance	12.46375
95%	169.1894	169.799	Skewness	-1.392987
99%	169.926	169.926	Kurtosis	5.127077

-> htgrp = 1

		Height (cm)		
	Percentiles	Smallest		
1%	170.0022	170.0022		
5%	170.18	170.0022		
10%	170.18	170.18	0bs	102
25%	171.45	170.18	Sum of Wgt.	102
50%	172.72		Mean	172.277
		Largest	Std. Dev.	1.338051
75%	173.482	173.99		
90%	173.99	174.1932	Variance	1.790381
95%	173.99	174.498	Skewness	1539058
99%	174.498	174.625	Kurtosis	1.892441

^{-&}gt; htgrp = 2

		Height	(cm)
1%	Percentiles 175.006	Smallest 175.006	

5% 10% 25%	175.26 175.26 175.895	175.006 175.26 175.26	Obs Sum of Wgt.	83 83
50%	176.53	Largest	Mean Std. Dev.	177.0533 1.355743
75%	177.8	179.07		
90%	179.07	179.07	Variance	1.838039
95%	179.07	179.07	Skewness	.1356182
99%	179.07	179.07	Kurtosis	1.75807

^{-&}gt; htgrp = 3

(cm)	Height
(Cill)	neignc

	Percentiles	Smallest		
1%	180.34	180.34		
5%	180.34	180.34		
10%	180.34	180.34	Obs	55
25%	180.34	180.34	Sum of Wgt.	55
50%	182.88		Mean	182.8292
		Largest	Std. Dev.	2.570328
75%	184.15	187.96		
90%	186.69	187.96	Variance	6.606587
95%	187.96	189.23	Skewness	1.049664
99%	190.5	190.5	Kurtosis	3.490751

Height (cm)

no observations

^{-&}gt; htgrp = .

II. Rate Ratios.

. stmh hieng

failure _d: chd
analysis time _t: (dox-origin)/365.25

origin: time dob enter on or after: time doe

Maximum likelihood estimate of the rate ratio comparing hieng==1 vs. hieng==0

RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	[95% Conf.	Interval]
0.520	4.84	0.0277	0.288	0.941

. stmh hieng,c(0,1)

failure _d: chd

analysis time _t: (dox-origin)/365.25 origin: time dob enter on or after: time doe

Maximum likelihood estimate of the rate ratio comparing hieng==0 vs. hieng==1

RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	[95% Conf.	Interval]
1.922	4.84	0.0277	1.063	3.474

In its simplest use, stmh estimates the ratio of the rates of failure for two categories of the explanatory variable (the first argument). Categories to be compared may be defined, as recode rules, in the compare option.

III. Exposure with more than two levels.

1. . egen eng3=cut(energy), at(1.5, 2.5, 3.0, 4.5) icodes

. tab eng3

eng3	Freq.	Percent	Cum.
0 1 2	75 150 112	22.26 44.51 33.23	22.26 66.77 100.00
Total	337	100.00	

. strate eng3, per(1000)

failure _d: chd
analysis time _t: (dox-origin)/365.25

origin: time dob enter on or after: time doe

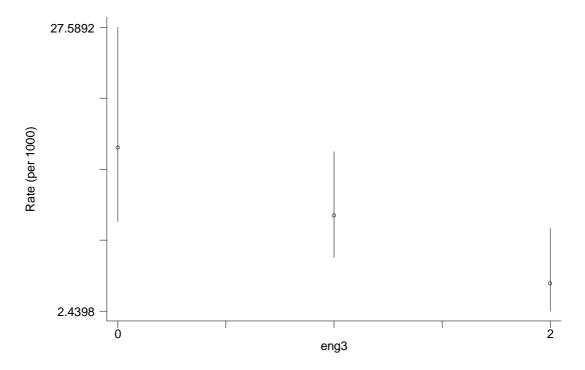
Estimated rates (per 1000) and lower/upper bounds of 95% confidence intervals $\,$

(337 records included in the analysis)

	eng3	D	У	Rate	Lower	Upper
	0	16 22	0.9466 2.0173	16.9020 10.9059	10.3547 7.1810	27.5892 16.5629
į	2	8	1.6398	4.8787	2.4398	9.7555

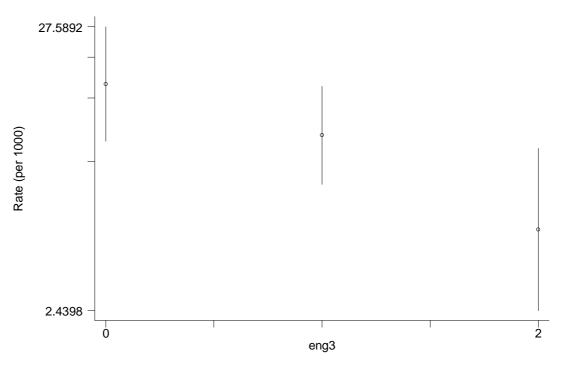
Here the rates of chd among the three levels of exposure are presented. Note that there is a decline in the actual rates as we go from the first to the last level.

. strate eng3, per(1000) graph



Note that with the option graph we can plot the actual rates. Therefore, this graph offers a graphical inspection of the actual difference in the rates.

. strate eng3, per(1000) graph ylog



With the additional option ylog we plot the log(rate). Note here that the differences we see are differences between the log(rates) and therefore we can inspect the log of the rate ratio (log(R1)-log(R2)=log(R1/R2)).

3. stmh eng3, c(1,0)

failure _d: chd

analysis time $_t$: (dox-origin)/365.25

origin: time dob enter on or after: time doe

Maximum likelihood estimate of the rate ratio comparing eng3==1 vs. eng3==0 RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	[95% Conf.	Interval]
0.645	1.81	0.1789	0.339	1.229

. stmh eng3, c(2,0)

failure _d: chd

analysis time _t: (dox-origin)/365.25

origin: time dob enter on or after: time doe

Maximum likelihood estimate of the rate ratio comparing eng3==2 vs. eng3==0 RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	[95% Conf.	<pre>Interval]</pre>
0.289	9.35	0.0022	0.124	0.674

4. . stmh eng3

failure _d: chd

analysis time _t: (dox-origin)/365.25

origin: time dob enter on or after: time doe

Score test for trend of rates with eng3 with an approximate estimate of the rate ratio for a one unit increase in eng3

RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	[95% Conf.	Interval]
0.548	8.98	0.0027	0.370	0.812

Here we deal with the exposure in its metric form. Therefore we will estimate an average effect of changing from the one level to the other. This is given by the RR = 0.548. It seems that the rate of CHD declines as we go from the lowest to the highest level of energy intake. This reduction is statistically significant.

5. . stmh energy

failure _d: chd
analysis time _t: (dox-origin)/365.25

origin: time dob enter on or after: time doe

Score test for trend of rates with energy with an approximate estimate of the rate ratio for a one unit increase in energy

RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	 [95% Conf.	Interval]
0.351	9.89	0.0017	0.183	0.674

Here we assess the effect of 1 unit (1 Mcal) increase in the actual value of energy intake. The rate ratio of 0.351 indicates a reduction with 1 Mcal increase in the energy intake. As expected, this is statistically significant also.

IV. Controlling for Confounding.

1. . stmh hieng job

failure _d: chd

analysis time _t: (dox-origin)/365.25

origin: time dob enter on or after: time doe

Mantel-Haenszel estimate of the rate ratio comparing hieng==1 vs. hieng==0 controlling for job

RR estimate, and lower and upper 95% confidence limits

RR	chi2	P>chi2	[95% Conf.	<pre>Interval]</pre>
0.525	4.71	0.0299	0.291	0.949

Here the aim is to explore the effect of energy intake on the rates of CHD, adjusting for any effect from the variable job. However, the previous form of the command stmh does not show the RRs within each category of job. We obtain this information with the following form of the command:

. stmh hieng, by(job)

Maximum likelihood estimate of the rate ratio
comparing hieng==1 vs. hieng==0
by job

RR estimate, and lower and upper 95% confidence limits

job	RR	Lower	Upper
0 1	0.41 0.66	0.12 0.23	1.36 1.89
2	0.52	0.21	1.27

Overall estimate controlling for job

RR	chi2	P>chi2	[95% Conf.	Interval]
0.525	4.71	0.0299	0.291	0.949

Approx test for unequal RRs (effect modification): chi2(2) = 0.33Pr>chi2 = 0.8468

What we get now is the effect of high vs. low energy intake on the rates of CHD within each category of job. We can see that these rates do not seem very different in between. We get a test for effect modification, i.e. a test for an interaction between job and hieng in the last lines of this output. The test is not statistically significant (p = 0.8468) but don't forget that it lacks power.

The <u>adjusted</u> estimated risk ratio for the effect of hieng on the rates of CHD is now 0.525, whereas the crude one was 0.520. Practically the adjustment did not affect the rate ratio. This is so because job does not affect the outcome (try this with stmh job).

2. . stmh hieng, by(job htgrp)

failure _d: chd

analysis time _t: (dox-origin)/365.25 origin: time dob enter on or after: time doe

Maximum likelihood estimate of the rate ratio comparing hieng==1 vs. hieng==0

by job htgrp

RR estimate, and lower and upper 95% confidence limits

job	htgrp	RR	Lower	Upper
0	150	0.69	0.10	4.88
0	170	0.41	0.04	3.95
0	175	0.22	0.02	2.15
1	150	0.59	0.15	2.37
1	170	1.09	0.18	6.52
2 2 2	150	0.41	0.08	2.26
	170	0.91	0.18	4.51
	175	0.50	0.11	2.22

Overall estimate controlling for job htgrp

RR	chi2	P>chi2	[95% Conf.	Interval]
0.569	3.48	0.0620	0.313	1.037

Approx test for unequal RRs (effect modification): chi2(7) = 1.85

Here we explore the effect of hieng whilst controlling for two confounders, job and htgrp. The strata are now defined by the cross-classification of the two possible confounders (i.e., job and htgrp). The rate ratios are estimated within each of these strata. As you can see, the high energy group does not affect statistically significantly the rate of CHD in any of the strata (all CIs include 1).

The MH estimate of 0.569 is relevant if there is not any interaction between the three variables. In terms of modelling, this is a model which fits the main effects of job, htgrp and hieng as explanatory variables and the rate of chd as the dependent variable. What we observe in the rates within strata is that the trend in the rates between the categories of htgrp within each of the levels of job is not the same. Perhaps the assumption of no interaction is not relevant...